

The Fundamental Role of General Orthonormal Bases in System Identification

Brett Ninness*

Håkan Hjalmarsson[†]

Fredrik Gustafsson[‡]

Abstract

The purpose of this paper is threefold. Firstly, it is to establish that contrary to what might be expected, the accuracy of well known and frequently used asymptotic variance results can depend on choices of fixed poles or zeros in the model structure. Secondly, it is to derive new variance expressions that can provide greatly improved accuracy while also making explicit the influence of any fixed poles or zeros. This is achieved by employing certain new results on generalised Fourier series and the asymptotic properties of Toeplitz-like matrices in such a way that the new variance expressions presented here encompass pre-existing ones as special cases. Via this latter analysis a new perspective emerges on recent work pertaining to the use of orthonormal basis structures in system identification. Namely, that orthonormal bases are much more than an implementational option offering improved numerical properties. In fact, they are an intrinsic part of estimation since, as shown here, orthonormal bases quantify the asymptotic variability of the estimates whether or not they are actually employed in calculating them.

Technical Report EE9739, Centre for Integrated Dynamics and Control (CIDAC),
Department of Electrical and Computer Engineering,
University of Newcastle, AUSTRALIA

1 Introduction

A result that has found great utility in practical applications of least-squares system identification methods [8, 21, 2] is that the sensitivity to measurement noise of the ensuing frequency response estimate $\hat{G}(e^{j\omega})$ may be quantified as [25, 22, 24, 50]

$$\text{Var}\{\hat{G}(e^{j\omega})\} \approx \frac{n}{N} \frac{\Phi_v(\omega)}{\Phi_u(\omega)} \quad (1)$$

*This work was supported by the Australian Research Council and the Centre for Integrated Dynamics and Control (CIDAC), University of Newcastle, Callaghan 2308, Australia. This author is with CIDAC and the Department of Electrical & Computer Engineering, University of Newcastle, Australia and can be contacted at email:brett@ee.newcastle.edu.au or FAX: +61 49 21 69 93. Part of this work was completed while visiting S3-Automatic Control, The Royal Institute of Technology, Stockholm, Sweden.

[†]Work partly completed while on leave at CIDAC from the Department of Sensors, Signals and Systems, The Royal Institute of Technology, S-100 44 Stockholm, Sweden. This author can be contacted at email:hakan.hjalmarsson@s3.e.kth.se or FAX: +46 8 790 7329

[‡]Work partly completed while on leave at CIDAC from the Department of Electrical Engineering, Linköping University, S-581 83, Linköping, Sweden. This author can be contacted at email:fredrik@isy.liu.se or FAX: +46 013 282622

where $\Phi_\nu(\omega)$ and $\Phi_u(\omega)$ are the measurement noise and input excitation spectral densities (respectively), N is the length of the available data record, and n is the order of the model $\widehat{G}(e^{j\omega})$.

In [25, 22, 24] an essential principle in deriving (1) is to establish that

$$\lim_{N, n \rightarrow \infty} \frac{N}{n} \text{Var}\{\widehat{G}(e^{j\omega})\} = \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} \quad (2)$$

and then assume that in fact (2) holds approximately for finite N and n , hence the result (1).

A main contribution of this paper is to show that in fact, for several commonly occurring situations, the approximation (1) can be quite misleading, even for large N and n . For example, problems are shown to occur when employing model structures with fixed denominator terms or when utilizing extensions of ARX structures with fixed noise model zeros.

Given these new observations, a more important purpose of this work is to establish how (1) should be modified so as to provide improved accuracy. In deriving these enhanced accuracy expressions, the third and final main theme of this paper emerges. Namely, to provide a new perspective on the role of orthonormal model structures in system identification.

To elaborate more on this final point, there has recently been significant activity on the system identification application of special classes of fixed denominator model structures that are constructed to be orthonormal; see for example [44, 45, 15, 37, 34] and the references therein. This latter work has been motivated by two main factors. Firstly, the achievement of improved numerical conditioning and secondly, the provision of efficient parameterisations that allow decreased variance error while still minimising bias error.

However, this second feature is not dependant on the orthonormal property of the model structure. This is so since the orthonormal structure is, under a linear parameter space transform, equivalent to any other equivalently flexible non-orthonormal structure with the same fixed poles. As well, since least squares methods are employed, the estimates depend linearly on the output measurements. Therefore, modulo numerical issues, *precisely* the same estimates of the system frequency response are obtained whether or not one goes to the effort of implementing an orthonormal structure, or a simpler structure wherein (for example) the denominator is fixed, and only the numerator co-efficients are estimated.

Given this observation, a key contribution of this paper is to discover a new role for orthonormal bases in a system identification context that is beyond their implementational advantage of improved numerical conditioning. This insight, which involves exposing that orthonormal bases play a far more fundamental role in system identification theory than has been previously suggested, is achieved by taking a different approach to that of previous work [44, 45, 15, 37, 34] where only model structures parameterised in terms of orthonormal bases have been considered. In contrast, it is shown here how orthonormal bases may also be used as an *analysis* tool to determine the properties of estimates derived with respect to *any* model structure.

In particular, this strategy achieves the afore-mentioned goal of improving the accuracy of (1) by replacing n in (1) with a function $K_n(\omega, \omega)$ which is the so-called ‘reproducing kernel’ associated with a particular n dimensional subspace and which depends on fixed poles or zeros. Note that even if such fixed components are not explicitly present in the model structure, they are implicitly there whenever the common practice of data pre-filtering is employed; see § 7.

As well, by utilizing the orthonormal basis formulation of [34], an explicit formula for $K_n(\omega, \omega)$ can be generated which clearly shows how $K_n(\omega, \omega)$, and hence $\text{Var}\{\widehat{G}(e^{j\omega})\}$, is affected by fixed model structure poles or zeros. For example, it will be seen that in the special case of all the poles/zeros fixed at the origin (FIR/ARX model structures), then $K_n(\omega, \omega) = n$ and hence (1) arises as a special case. However, if all the poles are not chosen at the origin, then inclusion of the $K_n(\omega, \omega)$ term in

place of n in (1) is shown here to smoothly extend the original analysis of [22, 25] and in so doing provide more accurate approximation of $\text{Var}\{\widehat{G}(e^{j\omega})\}$ over a wider range of circumstances.

The paper is structured as follows. In the following § 2 a short simulation example demonstrating the possible inaccuracy of the variance error approximation (1) is presented to motivate the ensuing analysis. The paper then provides a formal problem definition in § 3 and in so doing allows § 4 to theoretically examine and hence isolate the cause of the inaccuracy demonstrated numerically in § 2. This leads to the work of § 5 where it is shown how new techniques, based on new system theoretic results pertaining to generalised orthonormal bases, generalised Fourier convergence and generalised Toeplitz-like matrices may be employed to provide a variance error approximation which, in some circumstances (such as those illustrated in § 2), is a great improvement over the pre-existing one (1). In § 6 it is shown how the same new tools may be used to also quantify the bias error for fixed denominator model structures. By the end of § 6 the paper has exclusively focused on fixed denominator generalisations of the FIR structure, but in § 7 the scope is broadened to also consider ARX-like structures with fixed zeros in the noise model. In analogy with the preceding analysis, it is demonstrated that if all the noise model zeros are not fixed at the origin, then the asymptotic variance expressions derived in [22, 24] although prima facie applicable, are also liable to suffer from inaccuracy, the severity of which depends on the proportion of non-origin placed zeros. Motivated by this, it is demonstrated in § 7 how the same new techniques and results of previous sections can be brought to bear in the ARX modelling context so as to provide new approximations whose validity is not degraded by choosing noise model zeros away from the origin. Finally § 8 provides some concluding perspectives on the results and techniques of the paper.

There have been a number of contributions dealing with certain aspects of the problems considered in this paper, foremost of which are [25, 22, 44, 45, 15, 37, 10]. In the sequel it will be commented on how these foundations relate to the paper at hand.

2 Motivation

In the interests of clearly motivating the analysis of this paper consider the simple case of estimating, via least-squares, an $n = 12$ 'th order FIR model of the true system (the zero order hold equivalence is calculated using a sampling period of 1 second)

$$G(q) = \mathcal{ZOH} \left\{ \frac{1}{(s+1)(2s+1)} \right\} = \frac{0.1548q + 0.0939}{(q - 0.6065)(q - 0.3679)}$$

by observing 10000 samples of its input-output response when the former is a stationary Gaussian process with spectral density $\Phi_u(\omega) = 0.25/(1.25 - \cos \omega)$ and the latter is corrupted by zero mean Gaussian white noise of variance $\sigma^2 = 0.001$. In this case, since both n and N can reasonably be considered 'large', then the approximation (1) could be expected to hold. This can be checked by Monte-Carlo simulation over say, 500 input and noise realisations so as to estimate the variance $\text{Var}\{\widehat{G}(e^{j\omega})\}$ by its sample average, which can then be compared to the approximation (1). The results for just such an experiment are shown in Figure 1 with the agreement between (dashed-dot line) the expression (1) and the sample average (solid line) being excellent. Note that in this simulation (and in all the rest to follow in this section), the bias error is negligible, and hence the variance error represents the total error.

Now suppose, as may commonly occur in practice, that prior knowledge of the poles of $G(q)$ exists. Then in the interests of decreasing the bias error it makes sense to try to incorporate this prior knowledge in the estimation process by fixing some poles in the model near where it is believed

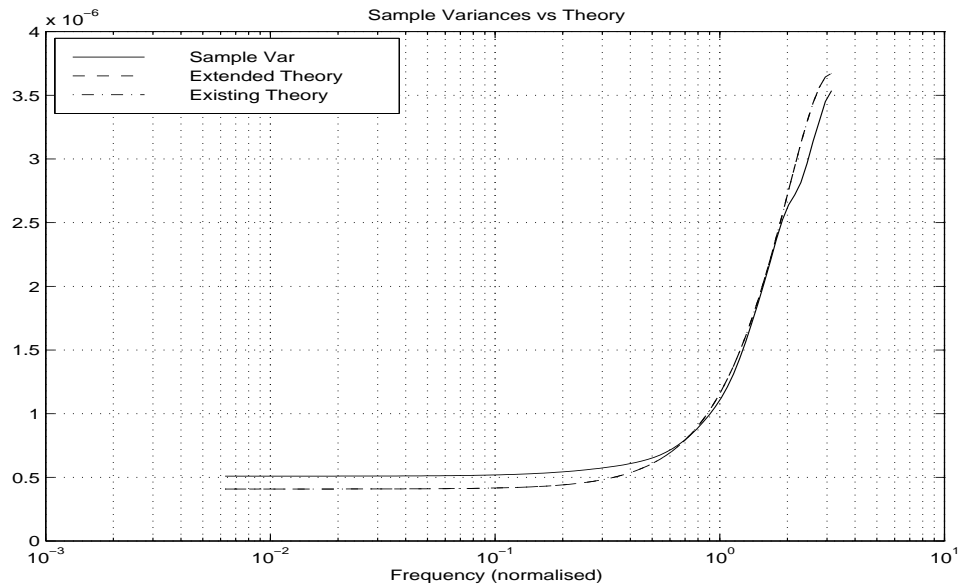


Figure 1: *FIR with all poles at origin. This is a comparison of a Monte–Carlo estimate of sample variability (solid line) with (dashed-dot line) the approximate expression (1). Note that this last line obscures a dashed line which is the new approximation (53). The obfuscation occurs because for the case of all poles at the origin the pre-existing approximation (1) and the new one (53) are identical.*

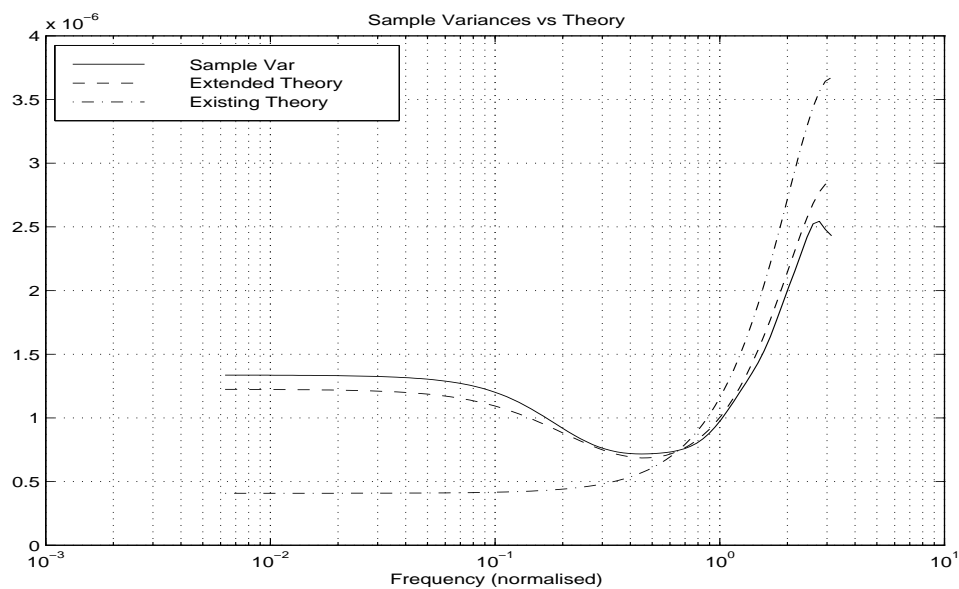


Figure 2: *FIR with 4 poles away from origin. This is a comparison of Monte–Carlo estimate of sample variability (solid line) with (dash-dot line) the approximate expression (2) and (dashed line) the new improved approximation (53) derived in this paper.*

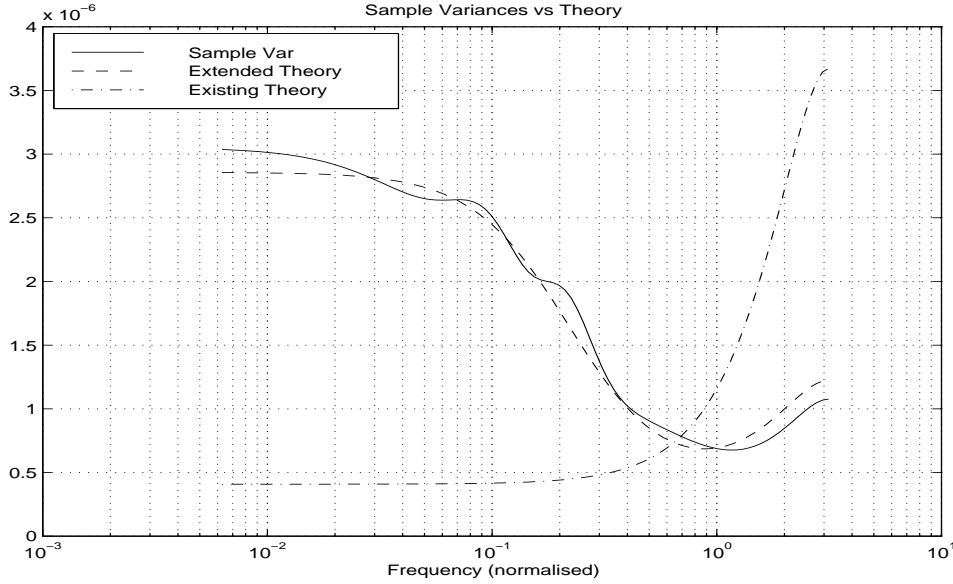


Figure 3: *FIR with 8 poles away from origin.*

the true poles lie [44, 15]. For example, suppose in the previous simulation it is believed that a dominant pole is near $z = 0.75$, so that guesses of, say $z = 0.7, 0.72, 0.78, 0.8$ are to be incorporated into the model structure. This can be implemented by simply pre-filtering the input by $F(q) = q^4 / (q - 0.7)(q - 0.72)(q - 0.78)(q - 0.8)$ before an FIR ‘numerator’ model $\hat{G}(q)$ is estimated, and then the complete system $\hat{G}'(q) = \hat{G}(q)F(q)$ may be taken as the fixed-pole estimate of the underlying dynamics.

Since the model order is still large ($n = 12$) then (1) should still provide a quantification of the variability of this new estimate by labeling the FIR estimate as \hat{G} and reasoning

$$\text{Var}\{\hat{G}'(e^{j\omega})\} = |F(e^{j\omega})|^2 \text{Var}\{\hat{G}(e^{j\omega})\} \approx \frac{n}{N} |F(e^{j\omega})|^2 \frac{\Phi_v(\omega)}{|F(e^{j\omega})|^2 \Phi_u(\omega)} = \frac{n}{N} \frac{\Phi_v(\omega)}{\Phi_u(\omega)} \quad (3)$$

which is unchanged from the normal FIR case. This unchanging nature is also reasonable, since the FIR case can be considered as already incorporating prior knowledge of system poles; namely poles near the origin.

Interestingly, when the expression (3) is compared to Monte–Carlo calculated sample variability as it is in Figure 2, then the agreement between the true variability (solid line) and approximation (3) (dash-dot line) is seen to be not nearly so good as is figure 1. Nevertheless, the expression (3) still provides useful information on the qualitative ‘high-pass’ nature of how the true variability changes with frequency. The dashed line near the solid one in figure 2 will be commented on in a moment.

Now suppose even more guesses of system poles are made, say at the locations $z = \{0.7, 0.72, 0.78, 0.8, 0.75, 0.85, 0.8$ with the sample variability again being compared to (3) in figure 3. In this case there is virtually no agreement (even qualitative) between true and predicted variability.

Clearly then, the well known approximation (1) can be quite misleading in situations where it would be expected to be reliable, and this apparently puzzling and practically important phenomenon is the motivation for this paper. In the work to follow, the conundrum just illustrated will be resolved by exposing certain hitherto unappreciated principles pertaining to when it is advisable to conclude (1) on the basis of (2). Furthermore, a new expression to extend (1) will be derived which is shown as

the dashed line in figures 1,2,3 and which (by consideration of those figures) can clearly offer greatly improved approximation of the true variability in certain circumstances.

3 Problem Formulation

Having provided certain motivating arguments, the remainder of the paper proceeds more formally. The problems studied in this paper are ones in which N point data records of an input sequence $\{u_t\}$ and output sequence $\{y_t\}$ of a linear time invariant system are available. It is assumed that this data is generated as follows

$$y_t = G(q)u_t + H(q)e_t.$$

Here $G(q)$ is an unknown transfer function describing, in terms of the forward shift operator q , the system dynamics that are to be identified by means of the observations $\{u_t\}$ and $\{y_t\}$. The output measurements $\{y_t\}$ are corrupted by a zero mean and stationary noise sequence $\nu_t = H(q)e_t$ where $H(q)$ is a stable and stably invertible and monic transfer function and $\{e_t\}$ is a zero mean white noise sequence with variance $\mathbf{E}\{e_t^2\} = \sigma^2 < \infty$ and with $\mathbf{E}\{|e_t|^{4+\delta}\} < \infty$ for some $\delta > 0$. The input $\{u_t\}$ is assumed to be quasi-stationary in the sense used by Ljung [20] (or, equivalently, to be amenable to the ‘generalised Harmonic analysis’ of Wiener [48]) so that it has an associated spectral density $\Phi_u(\omega) > 0$ which is assumed to be Lipschitz continuous of some order $\epsilon > 0$. The spectral density of the noise process $\{\nu_t\}$ is denoted as $\Phi_\nu(\omega) = \sigma^2 |H(e^{j\omega})|^2$ and is also assumed to be positive and Lipschitz continuous of order $\epsilon > 0$.

The identification schemes considered here are ones in which the prediction error framework [20, 43, 5] is used. This requires a model structure

$$y_t = G(q, \theta^n)u_t + H(q, \theta^n)e_t \quad (4)$$

parameterised by a vector $\theta^n \in \mathbf{R}^n$ to be employed which is of the form

$$G(q, \theta^n) = \sum_{k=1}^{\infty} g_k(\theta^n)q^{-k}, \quad H(q, \theta^n) = 1 + \sum_{k=1}^{\infty} h_k(\theta^n)q^{-k}.$$

This structure implies the following one step ahead predictor

$$\hat{y}_t(\theta^n) = [1 - H^{-1}(q, \theta^n)]y_t + H^{-1}(q, \theta^n)G(q, \theta^n)u_t \quad (5)$$

and associated prediction error

$$\varepsilon_t(\theta^n) \triangleq y_t - \hat{y}_t(\theta^n) \quad (6)$$

so that if the quadratic criterion

$$V_N(\theta^n) \triangleq \frac{1}{2N} \sum_{t=1}^N \varepsilon_t^2(\theta^n)$$

is employed, then based on the N point data observation, a least squares estimate $\hat{\theta}_N^n$ of θ^n may be found as

$$\hat{\theta}_N^n \triangleq \arg \min_{\theta^n \in \mathbf{R}^n} V_N(\theta^n). \quad (7)$$

The theory pertaining to the properties of such a method is very rich. Germane to this paper are the properties that [20, 43, 5]

$$\hat{\theta}_N^n \xrightarrow{\text{a.s.}} \theta_o^n \quad \text{as } N \rightarrow \infty$$

where with $\mathbf{E}\{\cdot\}$ denoting expectation over the underlying probability space that any random variables are defined on

$$\theta_{\circ}^n \triangleq \arg \min_{\theta^n \in \mathbf{R}^n} \lim_{N \rightarrow \infty} \mathbf{E} \{V_N(\theta^n)\}. \quad (8)$$

As well [23, 20]

$$\sqrt{N}(\hat{\theta}_N^n - \theta_{\circ}^n) \xrightarrow{\mathcal{D}} \mathcal{N}(0, P_n), \quad \text{as } N \rightarrow \infty$$

where

$$P_n \triangleq R_n^{-1} Q_n R_n^{-1} \quad (9)$$

and with the definition of the prediction error gradient $\psi_t(\theta^n)$ as

$$\psi_t(\theta^n) = \frac{d}{d\theta^n} \hat{y}_t(\theta^n)$$

then

$$\begin{aligned} R_n &\triangleq \lim_{N \rightarrow \infty} \mathbf{E} \left\{ \left. \frac{d^2 V_N(\theta)}{d\theta\theta^T} \right|_{\theta=\theta_{\circ}^n} \right\} = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{E} \{ \psi_t(\theta_{\circ}^n) \psi_t^T(\theta_{\circ}^n) \} - \\ &\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{E} \left\{ \varepsilon_t(\theta^n) \left(\frac{d\psi_t(\theta^n)}{d\theta^n} \right)^T \right\} \end{aligned} \quad (10)$$

and

$$\begin{aligned} Q_n &\triangleq \lim_{N \rightarrow \infty} N \mathbf{E} \left\{ \frac{dV_N(\theta_{\circ}^n)}{d\theta^n} \left(\frac{dV_N(\theta_{\circ}^n)}{d\theta^n} \right)^T \right\} \\ &= \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \sum_{\ell=1}^N \mathbf{E} \{ \psi_t(\theta_{\circ}^n) \psi_{\ell}(\theta_{\circ}^n)^T \varepsilon_t(\theta_{\circ}^n) \varepsilon_{\ell}(\theta_{\circ}^n) \}. \end{aligned} \quad (11)$$

A key contribution of [20, 22, 25] is to recognise that in applications, often the quantification of the parameter space properties of $\hat{\theta}_N^n$ are of secondary importance to their influence on the associated properties of $G(e^{j\omega}, \hat{\theta}_N^n)$ and $H(e^{j\omega}, \hat{\theta}_N^n)$. For the purposes of analysing the latter, it is expedient to define the composite transfer function

$$\Pi(q, \theta^n) \triangleq [G(q, \theta^n), H(q, \theta^n)] \quad (12)$$

and argue by Taylor expansion that

$$\begin{aligned} \tilde{\Pi}_N(\omega) &\triangleq \Pi(e^{j\omega}, \hat{\theta}_N^n) - \Pi(e^{j\omega}, \theta_{\circ}^n) \\ &= \left[\frac{d}{d\theta^n} \Pi(e^{j\omega}, \theta_{\circ}^n) \right]^T (\hat{\theta}_N^n - \theta_{\circ}^n) + o(\|\hat{\theta}_N^n - \theta_{\circ}^n\|) \end{aligned}$$

where here $\|\cdot\|$ denotes Euclidean norm so that with the notation $\Pi'(e^{j\omega}, \theta_{\circ}^n) \triangleq d\Pi(e^{j\omega}, \theta_{\circ}^n)/d\theta^n$ and with \cdot^* denoting ‘conjugate transpose’

$$\sqrt{N} \tilde{\Pi}_N(\omega) \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Lambda_n(\omega)) \quad (13)$$

where

$$\Lambda_n(\omega) = [\Pi'(e^{j\omega}, \theta_{\circ}^n)]^* P_n \Pi'(e^{j\omega}, \theta_{\circ}^n). \quad (14)$$

This suggests the approximation

$$N \mathbf{E} \left\{ \tilde{\Pi}_N(\omega) \tilde{\Pi}_N^*(\omega) \right\} \approx \Lambda_n(\omega).$$

Unfortunately, the evaluation of $\Lambda_n(\omega)$ is always too complicated to be useful. A key contribution of [22, 25] is to observe that in contrast to the intractability of $\Lambda_n(\omega)$, the limit

$$\lim_{N, n \rightarrow \infty} \frac{N}{n} \mathbf{E} \left\{ \tilde{\Pi}_N(\omega) \tilde{\Pi}_N^*(\omega) \right\} = \lim_{n \rightarrow \infty} \frac{1}{n} \Lambda_n(\omega) \triangleq \Lambda(\omega)$$

does have a simple and useful form; for example $\Lambda(\omega) = \Phi_\nu(\omega)/\Phi_u(\omega)$ for FIR model structures [25].

The suggestion of [22, 25] is to then assume that even for finite N and n it can reasonably be expected that $n^{-1}\Lambda_n(\omega)$ has approximately converged to the simple form $\Lambda(\omega)$ so that

$$\mathbf{E} \left\{ \tilde{\Pi}_N(\omega) \tilde{\Pi}_N^*(\omega) \right\} \approx \frac{n}{N} \Lambda(\omega) \quad (15)$$

is a good approximation. The validity of this strategy is argued via numerical example in [22, 25, 20], and indeed it has won widespread acceptance as a tool for analysing the performance of least squares estimation schemes; see for example [8, 21, 2, 12, 16].

However, § 2 has illustrated what seems to be the unappreciated fact that in certain circumstances the accuracy of (15) is not guaranteed even for what might be considered ‘large’ N or n . Instead, it is indicated in figures 1–3 that the presence of fixed model structure poles that are away from the origin may seriously degrade the accuracy of the approximation (1). The following section provides a theoretical analysis of this previously unrecognised phenomenon.

4 Fixed Denominator Model Structures

The observations of § 2 will, in the interests of generality, be addressed by considering the class of so-called ‘fixed denominator’ model structures. This terminology will refer to structures of the form (4) in which the poles of $G(q, \theta^n)$ are pre-specified and of which the archetypal example is

$$G(q, \theta^n) \triangleq D_n^{-1}(q) \sum_{k=0}^{n-1} \theta_k^n q^k, \quad H(q, \theta^n) = 1 \quad (16)$$

with

$$D_n(q) = \prod_{k=0}^{n-1} (q - \xi_k) \quad (17)$$

for some user chosen poles $\{\xi_k\} \in \mathbf{D} \triangleq \{z \in \mathbf{C} : |z| < 1\}$ where \mathbf{C} denotes the field of complex numbers. Although this is the quintessential form, the results to be presented here will apply to a much more general class of ‘fixed denominator’ model structures (54), (55) considered later.

Specialised ‘orthonormal’ versions of this structure have recently attracted significant research attention [44, 45, 15] where it has been suggested that although the choice $\xi_k = 0$ in (16) gives the common FIR structure, it is intuitively more reasonable to choose the poles $\{\xi_k\}$ according to prior knowledge so as to be close to the suspected true poles of $G(q)$.

In order to explain why when using these fixed denominator model structures, the prima facie applicable results of [25] lead to such a poor approximation illustrated in figure 3, it is necessary to

precisely examine the steps used in [25] to arrive at (1). For this purpose, first note that in the fixed denominator case the prediction error gradient ψ_t is given by

$$\psi_t(\theta^n) = \Gamma_n(q)D_n^{-1}(q)u_t = \Gamma_n(q)u'_t$$

where $u'_t \triangleq D_n^{-1}(q)u_t$ and

$$\Gamma_n(q) \triangleq [1, q, q^2, \dots, q^{n-1}]^T. \quad (18)$$

Therefore, $d\psi_t/d\theta^n = 0$ so that from (10)

$$R_n = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{E} \{ \psi_t(\theta^n) \psi_t^T(\theta^n) \}. \quad (19)$$

Defining for any positive function $f : [-\pi, \pi] \rightarrow (0, \infty)$ the $n \times n$ symmetric Toeplitz matrix $T_n(f)$ by (sometimes, in the interests of readability, ω will be used in place of $e^{j\omega}$)

$$T_n(f) \triangleq \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma_n(\omega) \Gamma_n^*(\omega) f(\omega) d\omega \quad (20)$$

then using Parseval's formula the matrix R_n may be written as [20]

$$R_n = T_n(\Phi_{u'}) \quad (21)$$

where clearly $\Phi_{u'} = |D_n(e^{j\omega})|^{-2} \Phi_u(\omega)$. It is also possible [22], for the case of $H(q) = 1$ to establish that provided the true system is in the model set then $Q_n = \sigma^2 R_n$ so that $P_n = \sigma^2 R_n^{-1} = \sigma^2 T_n^{-1}(\Phi_{u'})$. Furthermore, by employing well known results on the asymptotic behaviour of Toeplitz matrices [11, 47] it is also true that

$$T_n^{-1}(\Phi_{u'}) \approx T_n \left(\frac{1}{\Phi_{u'}} \right)$$

where the precise meaning of the \approx operator is that equality occurs in Hilbert–Schmidt weak matrix norm $|\cdot|$ (defined by $|A_n|^2 = n^{-1} \text{Trace}\{A_n^T A_n\}$) as $n \rightarrow \infty$.

Therefore, since in the fixed denominator case $G(q, \theta^n) = D_n^{-1}(q) \Gamma_n^T(q) \theta^n$ and $H(q, \theta^n) = 1$ so that $\Pi(q, \theta^n) = G(q, \theta^n)$ and hence $\Pi'(q, \theta^n) = D_n^{-1}(q) \Gamma_n(q)$ then from (14) and the preceding Toeplitz matrix approximation

$$\Lambda_n(\omega) \approx \frac{\sigma^2}{|D_n(e^{j\omega})|^2} \Gamma_n^*(\omega) T_n \left(\frac{1}{\Phi_{u'}} \right) \Gamma_n(\omega). \quad (22)$$

The final key ingredient facilitating the analysis in [25] is to recognise that

$$\frac{1}{n} \Gamma_n^*(\omega) T_n(f) \Gamma_n(\omega) = \sum_{k=-n}^n \left(1 - \frac{|k|}{n} \right) c_k e^{j\omega k} \quad (23)$$

where the $\{c_k\}$ are the Fourier co-efficients of $f(\omega)$ defined by

$$c_k \triangleq \frac{1}{2\pi} \int_{-\pi}^{\pi} f(\omega) e^{-j\omega k} d\omega$$

so that the left hand side of (23) is in fact a Cesàro mean Fourier reconstruction of f which is known [51], provided f is continuous, to converge uniformly to f on its domain and with increasing n .

The reasoning underlying the proofs of [25] for the special FIR case of $|D_n(e^{j\omega})| = |e^{j\omega n}| = 1$ then is that since by (23) the expression (22) is the n 'th order Cesàro mean Fourier reconstruction of $1/\Phi_{u'}(\omega)$, then for large n (22) should be approximately described by

$$\frac{N}{n} \text{Var}\{G(e^{j\omega}, \hat{\theta}_N^n)\} \approx \frac{1}{n} \Lambda_n(\omega) \approx \frac{\sigma^2}{|D_n(e^{j\omega})|^2} \frac{1}{\Phi_{u'}(\omega)} = \frac{\sigma^2}{\Phi_u(\omega)} \quad (24)$$

which is (1).

Now the mechanism leading to the possible inaccuracy of (1) is exposed. In order for (1) to be an accurate approximation, it is necessary that the n 'th order Cesàro mean Fourier reconstruction of $1/\Phi_{u'}(\omega)$ inherent in (22) has approximately converged. The size of n for which this may be expected to hold depends [51] on the smoothness of $\Phi_{u'}^{-1}(\omega) = |D_n(e^{j\omega})|^2/\Phi_u(\omega)$. For the special FIR case of $|D_n(e^{j\omega})| = 1$, the smoothness is fixed as the smoothness of $\Phi_u(\omega)$ so that the approximation can be expected to monotonically improve with increasing n . However, if poles are not chosen at the origin then the smoothness of $|D_n(e^{j\omega})|^2/\Phi_u(\omega)$ may decrease with increasing n so that convergence of (22) with increasing n need not occur.

Put another way, for poles of $D_n(q)$ not at the origin, the Fourier reconstruction inherent in (22) is with respect to co-efficients which are more accurately labeled as $\{c_k^n\}$ rather than $\{c_k\}$ since they depend on the model order n because they are the Fourier co-efficients of $|D_n|^2/\Phi_u$. In this case, as the length n of the Fourier reconstruction grows, the co-efficients may change if $|D_n(e^{j\omega})|$ changes, so there is no guarantee that approximate convergence has occurred for 'large' but finite n .

Therefore, the more poles that are chosen that are away from the origin, and hence the more dependant on n that $|D_n(e^{j\omega})|$ is, the less one should rely on (1) applying for finite n since the less likely it is that the underlying Fourier series has come close to convergence.

This is precisely the behaviour demonstrated in figures 1,2 and 3 of § 2.

5 Variance Error for Fixed Denominator Model Structure

Initially, this analysis of the utility of (1) for certain model structures may seem pessimistic since it indicates that much higher model orders n will be necessary before (1) can be used. A main contribution of this paper is to show that in fact this pessimism is unfounded, since as pre-cursed by the dashed lines of figures 2,3 it is possible to derive improved approximations that do not require increased model complexity in order to be accurate.

Central to achieving this goal is the consideration here of particular rational basis functions, hereafter denoted as the set $\{\mathcal{B}_k(q)\}$, and defined by ($\bar{\cdot}$ denotes complex conjugation of scalar quantities)

$$\mathcal{B}_0(q) \triangleq \frac{\sqrt{1-|\xi_0|^2}}{q-\xi_0}, \quad \mathcal{B}_k(q) \triangleq \frac{\sqrt{1-|\xi_k|^2}}{q-\xi_k} \prod_{m=0}^{k-1} \left(\frac{1-\bar{\xi}_m q}{q-\xi_m} \right), \quad k \geq 1. \quad (25)$$

They are orthonormal according to the inner product definition

$$\langle f, g \rangle \triangleq \frac{1}{2\pi} \int_{-\pi}^{\pi} f(e^{j\omega}) \overline{g(e^{j\omega})} d\omega = \frac{1}{2\pi j} \oint_{\mathbf{T}} f(z) \overline{g(z)} \frac{dz}{z} \quad (26)$$

where $\mathbf{T} \triangleq \{z \in \mathbf{C} : |z| = 1\}$; the orthonormality following easily using Cauchy's integral formula. These basis functions have been discussed in detail in [34] in the context of generalising certain other orthonormal bases existing in the system identification literature [44, 45, 37, 15].

One of the main contributions of this paper is to recognise that the asymptotic frequency domain properties of $G(e^{j\omega}, \hat{\theta}_N^n)$ are invariant to re-parameterisations of the model structure, and to therefore not attempt the scrutiny of (16) directly, but instead consider its reparameterised form

$$G(q, \theta^n) = \sum_{k=0}^{n-1} \theta_k^n \mathcal{B}_k(q), \quad H(q, \theta^n) = 1 \quad (27)$$

which, as will be demonstrated, is much more amenable to analysis than the original form (16).

Pivotal to the ensuing study employing this idea is the recognition of the importance of the the so-called ‘reproducing kernel’ $K_n(z, \mu)$ associated with the space $X_n \triangleq \text{Span}\{\mathcal{B}_0, \mathcal{B}_1, \dots, \mathcal{B}_{n-1}\}$ and which may be expressed as

$$K_n(z, \mu) = \sum_{k=1}^{n-1} \overline{\mathcal{B}_k(\mu)} \mathcal{B}_k(z), \quad z, \mu \in \mathbf{C}. \quad (28)$$

It derives its name from the property that for any $f \in X_n$ and $\mu \in \mathbf{C}$ such that $\mu \notin \{\xi_0, \dots, \xi_{n-1}\}$ then

$$f(\mu) = \langle f(z), K_n(z, \mu) \rangle \quad (29)$$

and by virtue of which $K_n(z, \mu)$ is unique [35]. In § 6 it will be shown how the property (29) can be exploited in order to quantify bias error, but to do so a more explicit expression for (28) is required. This has been derived in [35] where in analogy with equivalent formulae in the study of orthogonal polynomials [11, 41] it is termed a ‘Christoffel–Darboux’ formula:

$$K_n(z, \mu) = \frac{1 - \overline{\varphi_n(\mu)} \varphi_n(z)}{\overline{\mu} z - 1}, \quad z, \mu \in \mathbf{D}. \quad (30)$$

In this expression, $\varphi_n(z)$ is defined as the Blaschke product-like quantity

$$\varphi_0(z) = 1, \quad \varphi_n(z) \triangleq \prod_{k=0}^{n-1} \left(\frac{1 - \overline{\xi_k} z}{z - \xi_k} \right), \quad n \geq 1.$$

However, deferring the question of bias error until later, the purpose of this section is to quantify variance error, and for this purpose it is necessary to modify the original Toeplitz matrix formulation (20) to

$$M_n(f) \triangleq \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma_n(\omega) \Gamma_n^*(\omega) f(\omega) d\omega \quad (31)$$

which although formally identical to (31), is functionally quite different in that the underlying orthonormal basis is not fixed at the trigonometric one, but a generalisation obtained by re-defining $\Gamma_n(z)$ from (18) to

$$\Gamma_n(z) \triangleq [\mathcal{B}_0(z), \mathcal{B}_1(z), \dots, \mathcal{B}_{n-1}(z)]^T. \quad (32)$$

Matrices defined by (31), (32) will be called ‘generalised Toeplitz’ matrices, with the epithet deriving from the fact that if all the poles $\{\xi_k\}$ are chosen at the origin then $M_n(f) = T_n(f)$ is a bona-fide Toeplitz matrix, but otherwise it is not.

Pertinent to the employment of the matrices (31) in this paper is that with $\|\cdot\|$ denoting the matrix 2-norm [9] and as shown in [35] the following bounds apply

$$\|M_n(f)\| \leq \|f\|_{\infty}, \quad \|M_n^{-1}(f)\| \leq \|1/f\|_{\infty}. \quad (33)$$

Using the new notation (31), (32) the re-parameterised model structure (27) may be written as $G(q, \theta^n) = \Gamma_n^T(q)\theta^n$ so that $\psi_t(\theta^n) = \Gamma_n(q)$ and hence using (19) and Parseval's Formula R_n is a generalised Toeplitz matrix of the form

$$R_n = M_n(\Phi_u).$$

The matrix Q_n is more difficult to handle. To expand on this, note that with the formulation

$$z_k(\theta^n) \triangleq [G(q) - G(q, \theta^n)]u_k \quad (34)$$

then $\varepsilon_k(\theta^n) = z_k(\theta^n) + \nu_k$ and hence using the definition (11)

$$Q_n = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T [z_k(\theta^n) + \nu_k] [z_m(\theta^n) + \nu_m] \}. \quad (35)$$

If the data are collected in open loop so that $\{\psi_k\}$ and $\{\nu_k\}$ are uncorrelated then this can be re-written as

$$\begin{aligned} Q_n &= \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T \} \mathbf{E} \{ \nu_k \nu_m \} + \\ &\quad \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T z_k(\theta^n) z_m(\theta^n) \}. \end{aligned} \quad (36)$$

Considering only the first part of this expression, use of Lemma A.1 in the appendix permits its expression as

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T \} \mathbf{E} \{ \nu_k \nu_m \} = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma_n(\omega) \Gamma_n^*(\omega) \Phi_u(\omega) \Phi_\nu(\omega) d\omega. \quad (37)$$

This result (37) has been used in previous work [44, 45, 37, 34] related to the results of this section but despite its importance, to the knowledge of the authors, a proof of (37) has not previously appeared in any literature.

This still leaves the term containing $\{z_k(\theta^n)\}$ components to be dealt with. Clearly, from the definition (34) these terms are due to undermodelling, and if it so happens that $G(q)$ is in the model set, then they will be zero and hence (37) will completely describe Q_n . In the seminal work [25, 22], the assumption of $G(q)$ being in the model set for finite n was not made, and instead the effect of the $\{z_k(\theta^n)\}$ components was directly addressed. In subsequent work [44, 45, 37, 34] related to this paper it has been assumed that since eventually $n \rightarrow \infty$ is to be considered, then the effect of $z_k(\theta^n)$ will eventually be inconsequential and hence can be neglected from the beginning. However, since $z_k(\theta^n)$ enters into the formulation of $\Lambda_n(\omega)$ in such a complicated way, it is not obvious that its effect on $\Lambda_n(\omega)$ will tend to zero with increasing n even if $z_k(\theta^n)$ itself does.

As a result, this paper elects to in fact account for it by using (37) to write Q_n as

$$Q_n = M_n(\Phi_u \Phi_\nu) + \Delta_n$$

where Δ_n is the second term of (36) which, if neglected ab-initio as in [44, 45, 37], implies a presumption of $\lim_{n \rightarrow \infty} \|\Delta_n\| = 0$.

With these comments in hand, the analysis can proceed by noting that from (9)

$$P_n = M_n^{-1}(\Phi_u) M_n(\Phi_u \Phi_\nu) M_n^{-1}(\Phi_u) + M_n^{-1}(\Phi_u) \Delta_n M_n^{-1}(\Phi_u).$$

With the reparameterisation (27) then $G(q, \theta^n) = \Gamma_n^T(q)\theta^n$ with $\Gamma_n(q)$ being the generalised form (32) so that $\Pi'(q, \theta^n) = \Gamma_n(q)$ (the derivative with respect to $H(q)$ has been ignored since it is zero) and hence from (14) the quantity $\Lambda_n(\omega)$ quantifying variance error may be written as

$$\begin{aligned} \Lambda_n(\omega) &= \Gamma_n^*(\omega)M_n^{-1}(\Phi_u)M_n(\Phi_u\Phi_v)M_n^{-1}(\Phi_u)\Gamma_n(\omega) + \\ &\quad \Gamma_n^*(\omega)M_n^{-1}(\Phi_u)\Delta_nM_n^{-1}(\Phi_u)\Gamma_n(\omega). \end{aligned} \quad (38)$$

In order to simplify this for large n it is necessary to know something of the asymptotic algebraic structure of the generalised Toeplitz matrices in this expression, and this turns out to be a somewhat delicate question.

In related classical results [11, 47] pertaining to ‘classical’ symmetric Toeplitz matrices $T_n(f)$, it is shown that $|T_n(f)T_n(g) - T_n(fg)| \rightarrow 0$ as $n \rightarrow \infty$ where $|\cdot|$ denotes the Hilbert-Schmidt weak matrix norm defined earlier. In the generalised case an equivalent result cannot be shown to hold without overly restrictive assumptions being imposed [35]. Instead, it is necessary to be more subtle and recognise that in applications what is really important is how products $M_n(f)M_n(g)$ affect vectors only in the directions specified by $\Gamma_n(\omega)$.

This principle leads to the notion of asymptotic equivalence between two $n \times n$ matrices A_n and B_n with notation

$$A_n \sim B_n \quad \text{as } n \rightarrow \infty \quad (39)$$

if for all $\omega \in [-\pi, \pi]$

$$\lim_{n \rightarrow \infty} \frac{\Gamma_n^*(\omega)[A_n - B_n][A_n - B_n]^*\Gamma_n(\omega)}{K_n(\omega, \omega)} = 0. \quad (40)$$

The utility of this idea is that if $A_n \sim B_n$ as $n \rightarrow \infty$ then the asymptotic consideration of quadratic forms of A_n and B_n in $\Gamma_n(\omega)$ are identical. Specifically, for any $\omega, \lambda \in [-\pi, \pi]$ and using the Cauchy–Schwarz inequality

$$\begin{aligned} \left| \frac{\Gamma_n^*(\omega)A_n\Gamma_n(\lambda)}{K_n(\omega, \omega)} - \frac{\Gamma_n^*(\omega)B_n\Gamma_n(\lambda)}{K_n(\omega, \omega)} \right|^2 &= \left| \frac{\Gamma_n^*(\omega)[A_n - B_n]\Gamma_n(\lambda)}{K_n(\omega, \omega)} \right|^2 \\ &\leq \left| \frac{\Gamma_n^*(\omega)[A_n - B_n][A_n - B_n]^*\Gamma_n(\omega)}{K_n(\omega, \omega)} \right| \times \\ &\quad \left| \frac{\Gamma_n^*(\lambda)\Gamma_n(\lambda)}{K_n(\omega, \omega)} \right|. \end{aligned} \quad (41)$$

However, taking advantage of the explicit forms of the reproducing kernel made possible by (28) and (25)

$$\begin{aligned} \Gamma_n^*(\lambda)\Gamma_n(\lambda) = K_n(\lambda, \lambda) &= \sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\lambda})|^2 \\ &= \sum_{k=0}^{n-1} \frac{1 - |\xi_k|^2}{|e^{j\lambda} - \xi_k|^2} \end{aligned} \quad (42)$$

$$\leq \mu^2 \sum_{k=0}^{n-1} \frac{1 - |\xi_k|^2}{|e^{j\omega} - \xi_k|^2} = \mu^2 K_n(\omega, \omega) \quad (43)$$

where $\mu = \max_{\lambda, k} |e^{j\omega} - \xi_k| / |e^{j\lambda} - \xi_k|$ is some constant independent of n that is guaranteed to be finite provided all the poles $\{\xi_k\}$ are chosen in some closed subset of the open unit disk \mathbf{D} . Substituting

(43) in (41) shows that the latter tends to zero with increasing n thereby establishing the asymptotic equivalence of the quadratic forms. Note the key role played by the reproducing kernel $K_n(\omega, \omega)$.

Of fundamental importance in using these principles to clarify (38) are the results of [35] showing that if f and g are any (possibly complex valued) functions that are invertible and Lipschitz continuous of some order $\epsilon > 0$, and if all the poles $\{\xi_k\}$ are chosen in some closed subset of the open unit disc \mathbf{D} , then

$$M_n(f)M_n(g) \sim M_n(fg) \quad \text{as } n \rightarrow \infty, \quad (44)$$

$$M_n^{-1}(f) \sim M_n(1/f) \quad \text{as } n \rightarrow \infty. \quad (45)$$

Applying these results to (38) then provides the formulation

$$P_n \sim M_n \left(\frac{\Phi_\nu}{\Phi_u} \right) + M_n^{-1}(\Phi_u)\Delta_n M_n^{-1}(\Phi_u) \quad (46)$$

as $n \rightarrow \infty$ so that using the mechanism illustrated in (41)

$$\lim_{n \rightarrow \infty} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} = \lim_{n \rightarrow \infty} \frac{\Gamma_n^*(\omega)M_n(\Phi_\nu/\Phi_u)\Gamma_n(\omega)}{K_n(\omega, \omega)} + \delta_n(\omega) \quad (47)$$

where

$$\delta_n(\omega) \triangleq \frac{\Gamma_n^*(\omega)M_n^{-1}(\Phi_u)\Delta_n M_n^{-1}(\Phi_u)\Gamma_n(\omega)}{K_n(\omega, \omega)} \quad (48)$$

which using (33), the Cauchy–Schwarz inequality and the explicit form (28) can be bounded as

$$|\delta_n(\omega)| \leq \|1/\Phi_u\|_\infty^2 \|\Delta_n\|. \quad (49)$$

In order to further simplify (47) to some useful form, it is pivotal that a generalised analog of the Fourier convergence of (23) exists. Such a result is provided in [35] where it is shown that provided $\sum(1 - |\xi_k|) = \infty$ then for any possibly complex valued but continuous f

$$\lim_{n \rightarrow \infty} \frac{1}{K_n(\omega, \omega)} \Gamma_n^*(\omega)M_n(f)\Gamma_n(\lambda) = \begin{cases} f(\omega) & ; \omega = \lambda, \\ 0 & ; \omega \neq \lambda \end{cases} \quad (50)$$

Again observe the key role played by the reproducing kernel, and in particular note that use of the formulation (28),(25) with $\xi_k = 0$ gives $K_n(\omega, \omega) = n$ so that the left hand side of (50) becomes the Cesàro mean (23). Because of this, the result (50) can be seen to expand Fourier convergence analysis to the general orthonormal basis (25) which contains the classical trigonometric basis of (23) as a special case of $\xi_k = 0$. In fact, (50) is something more than this, since it also shows convergence to zero if $\lambda \neq \omega$. In the special trigonometric case this latter specialisation appears to have been first established in [13].

Applying (50) to (47) then provides the result that (assume for the moment that the limit $\lim_{n \rightarrow \infty} \delta_n(\omega)$ exists. This will be verified later.)

$$\lim_{n \rightarrow \infty} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} = \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} + \lim_{n \rightarrow \infty} \delta_n(\omega). \quad (51)$$

The dividend of considering the orthonormal reparameterisation (27) rather than the more natural fixed denominator form (16) is now revealed. Namely, in so doing (51) is derived from asymptotic analysis of an expression (47) which involves generalised Fourier analysis of an underlying function Φ_ν/Φ_u whose smoothness is invariant to model order, or fixed pole selection. This crucial feature will permit the derivation of the improved variance error approximation shown in figures 1–3.

Before further elucidating this point, the developments of this section are formally collected in the following theorem and corollary

Theorem 5.1. With $\widehat{\theta}_N^n$ calculated via (7) using the model structures (16), (27) or any other structure with the same fixed poles $\{\xi_k\}$ all chosen to lie within the open unit disk \mathbf{D} then with $K_n(\omega, \omega) \equiv K_n(e^{j\omega}, e^{j\omega})$ given by (28) and in the limit as $N \rightarrow \infty$

$$\sqrt{N} \begin{bmatrix} K_n(\omega, \omega) & 0 \\ 0 & K_n(\lambda, \lambda) \end{bmatrix}^{-1/2} \begin{pmatrix} G(e^{j\omega}, \widehat{\theta}_N^n) - G(e^{j\omega}, \theta_0^n) \\ G(e^{j\lambda}, \widehat{\theta}_N^n) - G(e^{j\lambda}, \theta_0^n) \end{pmatrix} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Sigma_n(\omega, \lambda))$$

where, provided the data is collected in open loop, all the poles $\{\xi_k\}$ are chosen in a closed subset of \mathbf{D} and $\Phi_u(\omega)$, $\Phi_v(\omega)$ are Lipschitz continuous of some order $\epsilon > 0$ then with the definition

$$\delta(\omega) \triangleq \lim_{n \rightarrow \infty} \delta_n(\omega)$$

where $\delta_n(\omega)$ is defined by (48), then

$$\lim_{n \rightarrow \infty} \Sigma_n(\omega, \lambda) = \begin{bmatrix} \frac{\Phi_v(\omega)}{\Phi_u(\omega)} & \delta'(\omega) \\ \delta'(\omega) & \frac{\Phi_v(\lambda)}{\Phi_u(\lambda)} \end{bmatrix} + \delta(\omega)I$$

where for some finite μ

$$|\widetilde{\delta}(\omega)| \leq \mu |\delta(\omega)|.$$

Proof. It has already been established in (13) and (14) that

$$\sqrt{N} \left[G(e^{j\omega}, \widehat{\theta}_N^n) - G(e^{j\omega}, \theta_0^n) \right] \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Lambda_n(\omega)) \text{ as } N \rightarrow \infty$$

where $\Lambda_n(\omega) = \Gamma_n^*(\omega) P_n \Gamma_n(\omega)$. An extension of this is that as $N \rightarrow \infty$

$$\sqrt{N} \begin{bmatrix} G(e^{j\omega}, \widehat{\theta}_N^n) - G(e^{j\omega}, \theta_0^n) \\ G(e^{j\lambda}, \widehat{\theta}_N^n) - G(e^{j\lambda}, \theta_0^n) \end{bmatrix} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \widetilde{\Sigma}_n(\omega, \lambda))$$

where

$$\widetilde{\Sigma}_n(\omega, \lambda) \triangleq \begin{bmatrix} \Gamma_n^*(\omega) P_n \Gamma_n(\omega) & \Gamma_n^*(\omega) P_n \Gamma_n(\lambda) \\ \Gamma_n^*(\lambda) P_n \Gamma_n(\omega) & \Gamma_n^*(\lambda) P_n \Gamma_n(\lambda) \end{bmatrix}$$

so that using the notation of (13), (14)

$$\sqrt{N} \begin{bmatrix} K_n(\omega, \omega) & 0 \\ 0 & K_n(\lambda, \lambda) \end{bmatrix}^{-1/2} \begin{pmatrix} G(e^{j\omega}, \widehat{\theta}_N^n) - G(e^{j\omega}, \theta_0^n) \\ G(e^{j\lambda}, \widehat{\theta}_N^n) - G(e^{j\lambda}, \theta_0^n) \end{pmatrix} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Sigma_n(\omega, \lambda))$$

where

$$\begin{aligned} \Sigma_n(\omega, \lambda) &= \begin{bmatrix} K_n(\omega, \omega) & 0 \\ 0 & K_n(\lambda, \lambda) \end{bmatrix}^{-1/2} \widetilde{\Sigma}_n(\omega, \lambda) \begin{bmatrix} K_n(\omega, \omega) & 0 \\ 0 & K_n(\lambda, \lambda) \end{bmatrix}^{-1/2}, \\ &= \begin{bmatrix} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} & \frac{\Gamma_n^*(\omega) P_n \Gamma_n(\lambda)}{\sqrt{K_n(\omega, \omega) K_n(\lambda, \lambda)}} \\ \frac{\Gamma_n^*(\lambda) P_n \Gamma_n(\omega)}{\sqrt{K_n(\omega, \omega) K_n(\lambda, \lambda)}} & \frac{\Lambda_n(\lambda)}{K_n(\lambda, \lambda)} \end{bmatrix}. \end{aligned}$$

It has just been established in (51) that

$$\lim_{n \rightarrow \infty} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} = \lim_{n \rightarrow \infty} \delta_n(\omega) + \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)}$$

which accounts for the asymptotic value of the diagonal terms of $\Sigma_n(\omega, \lambda)$. For the off-diagonal terms note that using the formulation (46) and the definition (48)

$$\tilde{\delta}(\omega) = \frac{\Gamma_n^*(\omega) P_n \Gamma_n(\lambda)}{\sqrt{K_n(\omega, \omega) K_n(\lambda, \lambda)}} = \frac{\Gamma_n^*(\omega) M_n(\Phi_\nu / \Phi_u) \Gamma_n(\lambda)}{\sqrt{K_n(\omega, \omega) K_n(\lambda, \lambda)}} + \frac{K_n(\omega, \omega) \delta_n(\omega)}{\sqrt{K_n(\omega, \omega) K_n(\lambda, \lambda)}}.$$

However, by the argument leading to (43), $K_n(\lambda, \lambda) \geq \mu^2 K_n(\omega, \omega)$ for some finite μ so that

$$\tilde{\delta}(\omega) \leq \mu \frac{\Gamma_n^*(\omega) P_n \Gamma_n(\lambda)}{K_n(\omega, \omega)} + \mu \delta(\omega).$$

Use of (50) then completes the proof. \square

The main use of this result will be to infer the nature of the variability of $G(e^{j\omega}, \hat{\theta}_N^n)$. However, as pointed out in [20], convergence in distribution guarantees nothing about mean-square convergence. Therefore, to be rigorous, it is necessary to separately consider this, but fortunately such examination is straightforward provided the assumptions on the noise process $\{e_t\}$ are somewhat strengthened.

Corollary 5.1. *Under the same conditions as the previous theorem, but with a strengthened requirement that $\mathbf{E}\{e_t^8\} < \infty$ then*

$$\lim_{n \rightarrow \infty} \lim_{N \rightarrow \infty} \frac{N}{K_n(\omega, \omega)} \mathbf{E}\{|G(e^{j\omega}) - G(e^{j\omega}, \theta_\circ^n)|^2\} = \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} + \delta(\omega) \quad (52)$$

where

$$\delta(\omega) \triangleq \lim_{n \rightarrow \infty} \delta_n(\omega)$$

with $\delta_n(\omega)$ defined by (48).

Proof. Follows from the previous Theorem using the methods in Appendix 9B of [20]. \square

By drawing on the precedent of [25, 22] in which the approximation (1) is derived from the asymptotic result (2), this paper uses the new asymptotic result (52) and the explicit formulae (28), (25) to suggest the approximation

$$\text{Var}\{G(e^{j\omega}, \hat{\theta}_N^n)\} \approx \frac{1}{N} \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} \sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2. \quad (53)$$

Note that in providing this approximation, it has been assumed as in [44, 45, 37] that the undermodelling induced component $\delta(\omega)$ is negligible, and this simplification will be commented upon later. However, before doing so, there are several important points pertaining to the approximation (53) that need to be highlighted.

Firstly, the expression (53) is shown as a dashed line in the figures 1–3 of § 2 which, being so close to the sample variability estimate of the true variability, illustrates that in situations where some fixed denominator poles are chosen away from the origin, the new expression (53) can represent an approximation that is significantly more accurate than the well known one (1). As already suggested, the reason for the improved approximation rests solely with the strategy of re-parameterising with

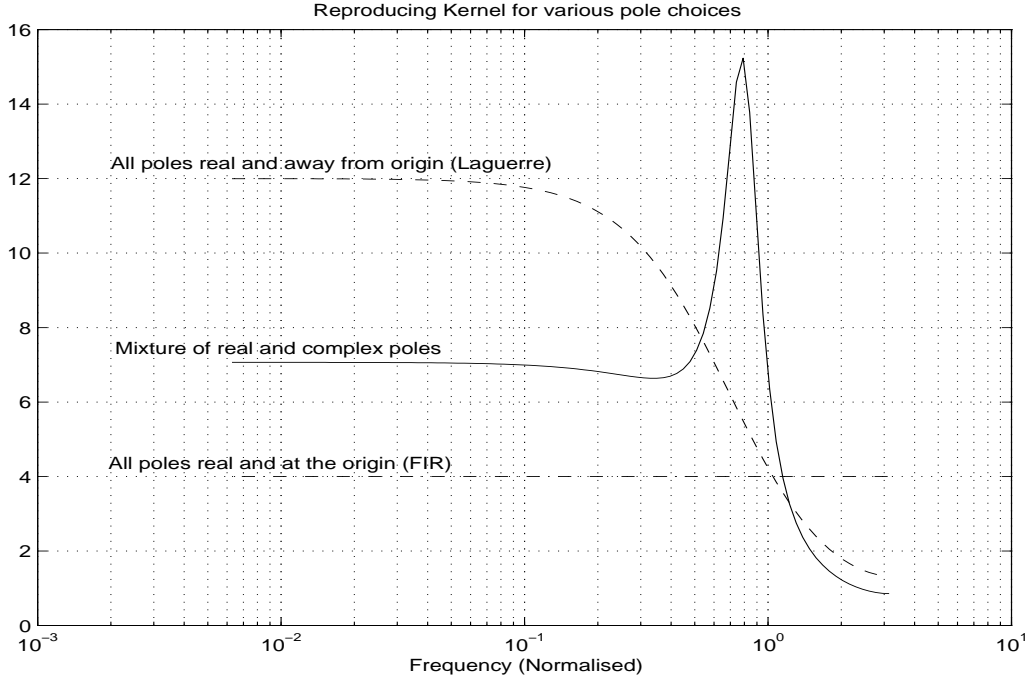


Figure 4: Plot, for various choices of $\{\xi_k\}$, of term $K_n(\omega, \omega) = \sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2$ which captures effect of pole choice $\{\xi_k\}$ on transfer function estimate sensitivity to measurement noise. Here $n = 4$.

respect to the orthonormal bases (25) in that by doing so, the asymptotic result (52) involves considering convergence of the generalised Fourier expansion (47) of a function Φ_ν/Φ_u whose smoothness is invariant to the model order, so that the degree of approximation in concluding (53) on the basis of (52) can be expected to monotonically improve with increasing model order n .

By using this novel reparameterisation approach, the essential feature imbuing the new expression (53) with greater accuracy than the pre-existing approximation (1) is that the influence of the fixed pole location on $\text{Var}\{G(e^{j\omega}, \hat{\theta}_N^n)\}$ is quantified by the reproducing kernel $K_n(\omega, \omega)$. See for example figure 4 where the expression $K_n(\omega, \omega)$ is plotted for a variety of choices of $\{\xi_k\}$. In particular, note that for all poles fixed at the origin, by the formulation (28),(25) then $K_n(\omega, \omega) = n$ so that in this special case of FIR modeling, (53) is identical to (1). However, the more poles that are not fixed at the origin, the more $K_n(\omega, \omega)$ will (being then frequency dependent) differ from n and hence the more the new approximation (53) will, in the interests of improved accuracy, be perturbed from the original approximation (1).

A point to note is that figure 4 indicates that the choice of poles $\{\xi_k\}$ could be conceived as a design variable for decreasing variance error at particular frequencies. However, not only might this incur an increased bias error if the $\{\xi_k\}$ are chosen far from the true poles (see the following section), figure 4 also indicates that decreased in variance error at one frequency are traded-off against increases at another frequency. In fact, it is a trivial consequence of orthonormality that the area (on a linear axis) underneath all the curves in figure 4 is the model order n .

Another point worth emphasising is that the results in Theorem 5.1 and Corollary 5.1 and hence in the approximation (53) apply for any model structure (of which (16) and (27) are special cases) of the form

$$G(q, \theta^n)u_t = \phi_t^T \theta^n \quad (54)$$

where

$$\phi_{t+1} = A\phi_t + Bu_t \quad (55)$$

with $A \in \mathbf{R}^{n \times n}$ and $B \in \mathbf{R}^{n \times 1}$ arbitrary but such that the eigenvalues of A are $\{\xi_0, \xi_1, \dots, \xi_{n-1}\}$. It seems quite unexpected that this role for the orthonormal bases $\{\mathcal{B}_k\}$ given by (25) should arise in such a fundamental manner in a problem that can via (54), (55) be ab-initio formulated with no orthonormality in its structure.

The role of orthonormal bases in system identification is therefore much deeper and more intrinsic than had previously been thought. This provides further motivation for their employment and study that is in addition to the numerical efficiency, and modelling efficiency features exposed in previous works [44, 37, 15, 34].

Indeed, previous work which has pioneered the use of orthonormal bases for system identification [44, 45, 37] has also considered variance expressions which are restricted versions of the form (53). For example, in [44] the use of Laguerre bases is analysed. However, this case may be encompassed in the framework of this paper by choosing all the poles the same and real valued as $\xi_k = \xi \in \mathbf{R}$ in which case (53) reduces to

$$\text{Var} \left\{ G(e^{j\omega}, \hat{\theta}_N^n) \right\} \approx \frac{n}{N} \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} \frac{1 - \xi^2}{|e^{j\omega} - \xi|^2} \quad (56)$$

which is the same as the result obtained in [44].

In [37, 15], which is the first work to consider the case of possibly different poles, they are still all restricted to be cyclically repeated from a finite set, say $\xi_k \in \{\xi_0, \dots, \xi_{p-1}\}$ and the result

$$\lim_{m, N \rightarrow \infty} \frac{N}{m K'_p(\omega, \omega)} \text{Var} \left\{ G(e^{j\omega}, \hat{\theta}_N^n) \right\} = \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} \quad (57)$$

is derived where $K'_p(\omega, \omega) = \sum_{k=0}^{p-1} |\mathcal{B}'_k(e^{j\omega})|^2$ with $\{\mathcal{B}'_k\}$ being the particular set of orthonormal bases considered in [37]. The quantity m is the number of *complete repetitions* of the *whole set* of bases $\{\mathcal{B}'_0, \dots, \mathcal{B}'_{p-1}\}$ in the model structure.

This latter point on complete repetitions of pole sets is of crucial importance since it has been demonstrated here that when inferring an approximation like (53) from an asymptotic result like (57) it is vital that convergence of a result like (57) can be reasonably expected to have occurred for finite m . However, if all the fixed poles in a model are chosen differently as in § 2 then $m = 1$. Therefore, since (57) is asymptotic in m , it cannot be argued that (57) can reasonably be expected to have converged in this case of $m = 1$, and hence the theoretical justification for an approximation like (53), which is derived on the basis of (57) in [37], is in these cases problematic.

On the other hand, using the methods of this paper, attention is focussed on the asymptotic nature of (47) which converges with increasing model order rather than pole set repetition count. In view of this (and with the precedent of the foundation work [44, 45, 37, 15] in mind) the new techniques proposed here are considered to provide a more theoretically sound argument for inferring a fixed order approximation from an asymptotic result.

Indeed, because of the fixed pole restrictions involved in previous work, the connections made here to general fixed denominator models of the form (54), (55) were never made and hence the fundamental relationship of orthonormal bases to situations which are not ab-initio formulated in terms of them has not been previously exposed.

A consequence of the pole restrictions imposed in the previous work [44, 45, 37] is that it forces the bases $\{\mathcal{B}'_n\}$ employed there to mimic the algebraic property $e^{j\omega n} e^{j\omega m} = e^{j\omega(m+n)}$ of the classical

trigonometric basis. This allows an elegant path to variance results to be followed wherein the problem is reduced to an equivalent FIR problem by establishing, via a bilinear or multilinear ('Hambo') transform, an algebra isomorphism between $\{\mathcal{B}'_n\}$ and $e^{j\omega n}$. The original results of [25] are then mapped through this same isomorphism in order to generalise them [44, 45, 37].

As is evident, the strategy of this paper is completely different and, apart from avoiding pole restrictions, it has the dividend of illustrating new analysis methods employing recently derived results [35] pertaining to generalised Toeplitz matrices and generalised Fourier analysis. These new techniques are considered to be of interest in their own right, as they would appear to have potential for application beyond the particular system identification setting considered here.

The final advantage of the methods pursued here are that they permit an explicit formulation (28),(25) for the term $K_n(\omega, \omega)$ revealing how fixed poles affect variability; in previous work it has been implicit in a state space construction [37]. Indeed, by the recognition here that this term $K_n(\omega, \omega)$ affecting variability is a reproducing kernel, and hence unique, a closed form expression for the equivalent term in [37] is provided for the first time.

It remains to comment on the undermodelling induced $\delta_n(\omega)$ term in the derivation of (53). Firstly, note that with its omission $\tilde{\delta}(\omega) = 0$ in Theorem 5.1 so that the frequency response estimates at different frequencies are asymptotically uncorrelated. Note also, that it can be argued that disregarding $\delta_n(\omega)$ is reasonable, since if one assumes for the sake of simplicity that $\mathbf{E}\{\psi_k \psi_m^T z_k(\theta_\circ^n) z_m(\theta_\circ^n)\} = \mathbf{E}\{\psi_k \psi_m^T\} \mathbf{E}\{z_k(\theta_\circ^n) z_m(\theta_\circ^n)\}$ then using Lemma A.1 and assuming $\{z_k(\theta_\circ^n)\}$ is a stationary stochastic process

$$\Delta_n = M_n (|G - G(\theta_\circ^n)|^2 \Phi_u^2)$$

so that use of (33) provides

$$\|\Delta_n\|_2 \leq \|G(e^{j\omega}) - G(e^{j\omega}, \theta_\circ^n)\|_\infty \|\Phi_u(\omega)\|_2^2.$$

In this case, using (49)

$$|\delta_n(\omega)| \leq \|G(e^{j\omega}) - G(e^{j\omega}, \theta_\circ^n)\|_\infty \|\Phi_u\|_\infty \|1/\Phi_u\|_\infty.$$

Since it will be proved in lemma 6.1 that the right hand side of this inequality tends to zero, theoretical justification which has not been provided in some previous works of neglecting the $\delta_n(\omega)$ term is provided.

In fact, this can be made more rigorous in such a way to even address the closed loop case, but at the cost of more restrictive assumptions.

Theorem 5.2. *Under the restriction that both $\{u_t\}$ and $\{e_t\}$ are Gaussian distributed, zero mean stationary stochastic processes, then the results of Theorem 5.1 and Corollary 5.1 hold even for closed loop data collection and with the $\delta(\omega)$ term guaranteed to be identically zero provided that the substitution*

$$\Phi_\nu(\omega) \mapsto \Phi_\varepsilon(\omega, \theta_\circ^n)$$

is made in the statements of Theorem 5.1 and Corollary 5.1. Here $\Phi_\varepsilon(\omega, \theta_\circ^n)$ is to be interpreted as the spectral density of the prediction error residual sequence $\{\varepsilon_t(\theta_\circ^n)\}$ defined in (6) evaluated at $\theta^n = \theta_\circ^n$, the latter being defined in (8).

Proof. The assumption of open loop data collection only manifests itself in the proofs of Theorem 5.1 and Corollary 5.1 by allowing the splitting of (35) into two terms in (36), the first of which can have lemma A.1 applied to it to conclude that $Q_n \sim M_n(\Phi_u \Phi_\nu) + \Delta_n$ as $n \rightarrow \infty$ where Δ_n accounts for the second, undermodelling induced term in (36). However, under Gaussian assumptions, and

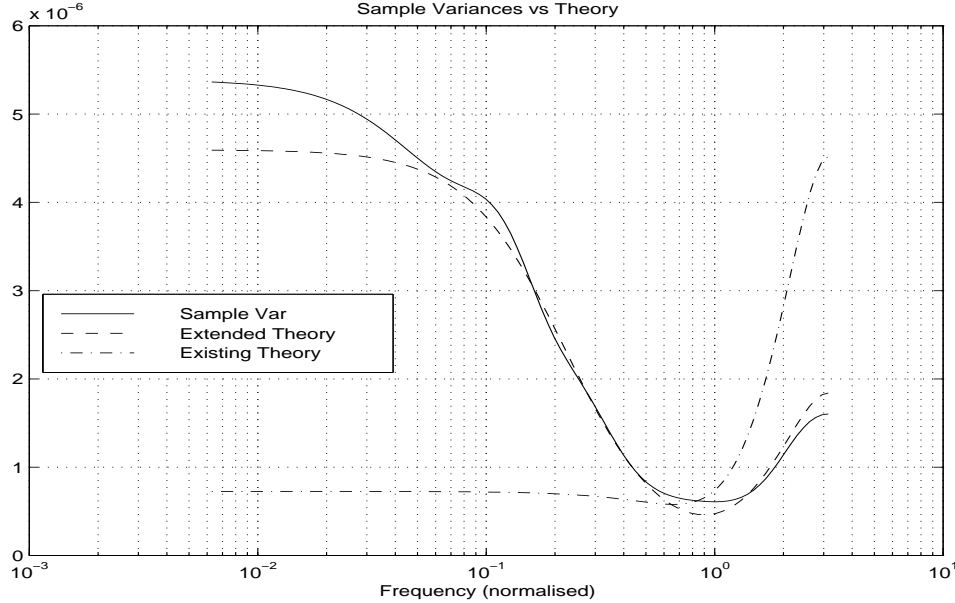


Figure 5: *FIR with 8 poles away from origin: Closed loop data collection and a non-Gaussian (uniformly) distributed process.*

regardless of whether or not the data is collected in open loop, then instead of this strategy lemma A.2 may be applied directly to establish that $Q_n \sim M_n(\Phi_u(\omega)\Phi_\nu(\omega, \theta_o^n))$ as $n \rightarrow \infty$ (with $\Delta_n \sim 0$). The rest of the proofs of Theorem 5.1 and Corollary 5.1 then go through unchanged. \square

Therefore, in the Gaussian case, the new approximation (53), which contains the pre-existing one (1) as a special FIR case, is shown theoretically to hold for closed loop data collection. This is the first time a result like (53) has been established for this case of ‘fixed denominator’ modelling, all the preceding work [25, 44, 45, 37] either applying only for open loop situations, or when considering closed-loop data [22], applies only for the true noise model being $H(q) = 1$.

To illustrate the efficacy of Theorem 5.2, consider the simulation conditions of § 2 altered so that the closed loop scenario of a proportional controller of gain $K = 1$ is involved and that furthermore the reference signal $\{r_t\}$ has spectrum $\Phi_r(\omega) = 0.25/(1.25 - \cos \omega)$ and the measurement noise has spectrum $\Phi_\nu(\omega) = (1.36 - 1.2 \cos \omega)/(1.16 - 0.8 \cos \omega)$. The sample average and theoretical variabilities given by (1) and (53) for this situation where 8 poles in the model structure are not at the origin are shown in figure 5. As in § 2, the new approximation (53) shown as the dashed line provides a more informative approximation to the true variability (solid line) than does the pre-existing approximation (1). Note that in using Theorem 5.2 to form the (dashed line) approximation in figure 5, it was assumed that (since the model structure was considered ‘rich’ and there were no common parameters between dynamics and noise model) there was negligible bias error, and hence that $\Phi_{\varepsilon(\theta_o^n)}(\omega) \approx \Phi_\nu(\omega)$.

An interesting point is that the simulation producing figure 5 specifically violated the Gaussian assumptions that Theorem 5.2 was derived under by using uniformly distributed random processes. This robustness of the result to violations of the assumptions it was derived under is considered encouraging from the viewpoint of its practical utility.

6 Bias Error with Fixed Denominator Model Structure

Having quantified the variance error involved with using the structure (16), the paper now turns to the question of quantifying the bias error, again by a strategy of considering the equivalent but more tractable orthonormal re-parameterisation (27). First, a key motivation for a model structure such as (16) is the intuitive belief that for fixed n , any undermodelling associated with lack of model structure richness can be decreased by choosing poles $\{\xi_k\}$ in $D_n(z)$ close to the true poles $\{\gamma_k\}$ in the underlying true system $G(z)$. The following Theorem justifies this intuition by use of the Christoffel–Darboux formula (30).

Theorem 6.1. *Suppose $G(z)$ has partial fraction expansion*

$$G(z) = \sum_{i=0}^{m-1} \frac{\alpha_i}{z - \gamma_i}$$

where all the poles satisfy $|\gamma_i| < 1$. Put $\widehat{G}_n(z)$ as the best L_2 approximation to $G(z)$ with respect to the n basis functions $\{\mathcal{B}_0, \mathcal{B}_1, \dots, \mathcal{B}_{n-1}\}$. Then

$$|G(e^{j\omega}) - \widehat{G}_n(e^{j\omega})| \leq \sum_{i=0}^{m-1} \left| \frac{\alpha_i}{e^{j\omega} - \gamma_i} \right| \prod_{k=0}^{n-1} \left| \frac{\gamma_i - \xi_k}{1 - \xi_k \gamma_i} \right|. \quad (58)$$

Proof. Take $\mu \in \mathbf{D}$. Then since $G(1/z)$ is analytic in \mathbf{D} the use of Cauchy's Integral Theorem provides

$$G(1/\mu) = \frac{1}{2\pi j} \oint_{\mathbf{T}} \frac{G(1/z)}{z - \mu} dz.$$

Also, by use of the defining property (29) of the reproducing kernel

$$\widehat{G}_n(1/\mu) = \sum_{k=0}^{n-1} \mathcal{B}_k(1/\mu) \frac{1}{2\pi j} \oint_{\mathbf{T}} G(t) \overline{\mathcal{B}_k(t)} \frac{dt}{t}.$$

Using a change of variable $z = t^{-1}$ then leads to

$$\widehat{G}_n(1/\mu) = \frac{1}{2\pi j} \oint_{\mathbf{T}} G(1/z) \left(\sum_{k=0}^{n-1} \mathcal{B}_k(1/\mu) \overline{\mathcal{B}_k(1/z)} \right) \frac{dz}{z},$$

so that using the Christoffel–Darboux Formula (30)

$$\begin{aligned} G(1/\mu) - \widehat{G}_n(1/\mu) &= \frac{\mu \varphi_n(1/\mu)}{2\pi j} \oint_{\mathbf{T}} \frac{G(1/z) \overline{\varphi_n(1/z)}}{z - \mu} \frac{dz}{z} \\ &= \frac{\mu \varphi_n(1/\mu)}{2\pi j} \oint_{\mathbf{T}} \frac{G(z) \overline{\varphi_n(z)}}{1 - \mu z} dz. \end{aligned} \quad (59)$$

Then by using (59) and Cauchy's Residue Theorem

$$\begin{aligned} G(1/\mu) - \widehat{G}_n(1/\mu) &= \frac{\mu \varphi_n(1/\mu)}{2\pi j} \sum_{i=0}^{m-1} \oint_{\mathbf{T}} \left(\frac{\alpha_i}{z - \gamma_i} \right) \left(\frac{1}{1 - \mu z} \right) \prod_{k=0}^{n-1} \left(\frac{z - \xi_k}{1 - \xi_k z} \right) dz, \\ &= \mu \varphi_n(1/\mu) \sum_{i=0}^{m-1} \left(\frac{\alpha_i}{1 - \mu \gamma_i} \right) \prod_{k=0}^{n-1} \left(\frac{\gamma_i - \xi_k}{1 - \xi_k \gamma_i} \right). \end{aligned}$$

Writing $\mu = re^{-j\omega}$ with $|r| < 1$, letting $r \rightarrow 1$ and using the triangle inequality then gives the result. \square

Note that this bound is tight in the sense that if perfect knowledge of the poles of $G(z)$ is available (so that $\forall i, \exists k : \xi_k = \gamma_i$) then the upper bound is zero. The interpretation of the result is that the convergence of the expansion (27) to the underlying system $G(q)$ can be very much faster than that of the special cases of FIR, Laguerre or two-parameter Kautz expansions if reasonable guesses $\{\xi_k\}$ of the poles of $G(q)$ can be made. This is so since

$$\frac{z - \xi_k}{1 - \xi_k z} \quad (60)$$

is analytic on \mathbf{D} and of modulus 1 on the boundary \mathbf{T} of \mathbf{D} . Therefore, by the maximum modulus Theorem, the factor (60) with $z = \gamma_i$ that appears n times in the bound (58) is of modulus less than one, and hence the bound (58) decreases geometrically with model order n ; the geometric rate will depend on the error term $|\gamma_i - \xi_i|$ in the guess ξ_i in the true pole position γ_i .

A bound similar to that of Theorem 6.1 is given, using completely different methods related to the new theory of ‘Hambo’ transforms in [15]. However, it is less explicit since it contains a constant scaling factor for which the only information available is that it is finite.

The significance of Theorem 6.1 in the context of quantifying the undermodelling induced error in an estimated frequency response is realised by noting that

$$\theta_{\circ}^n = \arg \min_{\theta^n \in \mathbf{R}^n} \left\{ \int_{-\pi}^{\pi} |G(e^{j\omega}) - G(e^{j\omega}, \theta^n)|^2 \Phi_u(\omega) d\omega \right\} \quad (61)$$

so that for $\{u_t\}$ white ($\Phi_u(\omega) = \text{constant}$), the expression (58) is a quantification of the asymptotic undermodelling induced estimation error.

Although (58) appears to be the most explicit statement that can be made about undermodelling induced error, it is possible to draw further conclusions that apply also when $\Phi_u(\omega)$ is not white. For example, it is shown in [31] that the frequency response magnitude is on average (over frequency) under-estimated in the sense that

$$\int_{-\pi}^{\pi} |G(e^{j\omega}, \theta_{\circ}^n)|^2 \Phi_u(\omega) d\omega \leq \int_{-\pi}^{\pi} |G(e^{j\omega})|^2 \Phi_u(\omega) d\omega.$$

The equation (61) may also be used to bound the bias error in estimating $G(e^{j\omega})$ by noting that since, under the condition $\sum(1 - |\xi_k|) = \infty$ the $\{\mathcal{B}_k\}$ are complete in $H_2(\mathbf{T})$, then it may be expanded as

$$G(e^{j\omega}) = \sum_{\ell=0}^{n-1} \theta_{\ell}^n \mathcal{B}_{\ell}(e^{j\omega}) + \sum_{\ell=n}^{\infty} \kappa_{\ell} \mathcal{B}_{\ell}(e^{j\omega}). \quad (62)$$

Using this notation, the following bias error bound can be obtained.

Lemma 6.1. *Suppose that $G(z) \in H_2(\mathbf{T})$ so that the co-efficients κ_{ℓ} in (62) are well defined. Then*

$$|G(e^{j\omega}) - G(e^{j\omega}, \theta_{\circ}^n)| \leq \left(n \frac{\max_{\omega} \Phi_u(\omega)}{\min_{\omega} \Phi_u(\omega)} + 1 \right) \max_{0 \leq k < n} |\mathcal{B}_k(\omega)| \sum_{\ell=n}^{\infty} |\kappa_{\ell}|. \quad w.p.I.$$

Proof. The method of proof is identical to that employed by Wahlberg in [44]. Expression (61) characterises θ_{\circ}^n as the solution to

$$R_n \theta_{\circ}^n = \mathbf{E} \{ \phi_k y_k \}.$$

Using this and defining the sequence $\{z_k\}$ as

$$z_k = \sum_{\ell=n}^{\infty} \kappa_{\ell} \mathcal{B}_{\ell}(q) u_k$$

gives

$$\|\theta^n - \theta_{\circ}^n\|_2 \leq \|R_n^{-1}\|_2 \|\mathbf{E} \{ \phi_k z_k \} \|_2.$$

But by Parseval's Theorem

$$[\mathbf{E} \{ \phi_k z_k \}]_m = \frac{1}{2\pi} \sum_{\ell=n}^{\infty} \kappa_{\ell} \int_{-\pi}^{\pi} \mathcal{B}_m(e^{j\omega}) \overline{\mathcal{B}_{\ell}(e^{j\omega})} \Phi_u(\omega) d\omega.$$

Now, using the triangle inequality, the Cauchy-Schwarz inequality and using (33) to bound $\|R_n^{-1}\|$ provides

$$\|\theta^n - \theta_{\circ}^n\|_1 \leq n \frac{\max_{\omega} \Phi_u(\omega)}{\min_{\omega} \Phi_u(\omega)} \sum_{\ell=n}^{\infty} |\kappa_{\ell}|.$$

But from (62)

$$|G(e^{j\omega}) - G(e^{j\omega}, \theta_{\circ}^n)| \leq \max_{0 \leq k < n} |\mathcal{B}_k(e^{j\omega})| \left(\|\theta^n - \theta_{\circ}^n\|_1 + \sum_{\ell=n}^{\infty} |\kappa_{\ell}| \right).$$

□

As mentioned for the particular case of Laguerre models studied in [44], this indicates that as $n \rightarrow \infty$ then θ_{\circ}^n will converge to the expansion co-efficients given in (62) and hence the frequency response estimate will converge to the true one provided the expansion co-efficients in (62) decrease sufficiently quickly that

$$\lim_{n \rightarrow \infty} n \sum_{\ell=n}^{\infty} |\kappa_{\ell}| = 0.$$

This is satisfied for example if $G(z)$ is stable and finite dimensional. In this case a simple argument using Cauchy's Residue Theorem shows that there exists a $|\eta| < 1$ and $K < \infty$ such that $|\kappa_{\ell}| < K\eta^{\ell}$.

7 ARX Type Model Structures

Although the fixed denominator model structure (16) and its generalisation (54),(55) have many practical advantages such as relevance to 'identification for control' ideas [1, 40, 27, 33] and the requirement of only simple and robust numerical procedures for the calculation of $\hat{\theta}_N^n$ [44, 37], they suffer from the drawback of relying on prior-knowledge of pole location.

A common strategy for avoiding this drawback is to estimate the pole locations of $G(q)$ while still involving a predictor that is linear in θ^n so that the advantage of simple numerical requirement for finding $\hat{\theta}_N^n$ is retained. This is done by employing the model structure

$$G(q, \theta^n) = \frac{B(q, \theta^n)}{A(q, \theta^n)}, \quad H(q, \theta^n) = \frac{D_n(q)}{A(q, \theta^n)} \quad (63)$$

where

$$\begin{aligned} A(q) &= a_0 + a_1q + a_2q^2 + \cdots + a_{n-1}q^{n-1} + q^n, \\ B(q) &= b_0 + b_1q + b_2q^2 + \cdots + b_{n-1}q^{n-1} \end{aligned}$$

with

$$\theta^n = [a_0, b_0, a_1, b_1, \cdots, a_{n-1}, b_{n-1}]^T$$

being the vector of parameters to be estimated and $D_n(q)$ is as previously defined in (17).

Here the dynamics and noise model share parameters in θ^n . As is well known, this can lead to bias if the model structure is not rich enough [46, 20]. The motivation for including the $D_n(q)$ term in the noise model is to avoid this bias by allowing $H(q, \theta^n) \approx H(q)$ for some θ^n while simultaneously, through zeros of $B(q, \theta^n)$ cancelling parts of $A(q, \theta^n)$ that pertain only to $H(q, \theta^n)$, achieving sufficient flexibility for $G(q, \theta^n) = G(q)$.

Most commonly the model structure (63) appears with the choice $\xi_k = 0$ in which case it is known as the ‘equation error’ or sometimes ‘ARX’ model structure and for which the analysis of Ljung [22] provides, using the definition (12), the well known result (which holds, when at least asymptotically in n , the true system is in the model set)

$$\text{Cov} \left\{ \Pi(e^{j\omega}, \hat{\theta}_N^n) \right\} \approx \frac{n}{N} \sigma^2 |H(e^{j\omega}, \theta^n)|^2 \Phi_\zeta^{-1}(e^{j\omega}, \theta^n). \quad (64)$$

Here

$$\Phi_\zeta(\omega, \theta^n) \triangleq \begin{bmatrix} \Phi_u(\omega) & \Phi_{u\varepsilon(\theta^n)}^*(\omega) \\ \Phi_{u\varepsilon(\theta^n)}(\omega) & \Phi_{\varepsilon(\theta^n)}(\omega) \end{bmatrix}$$

is the spectral density of the two dimensional vector signal $\zeta_t(\theta^n)$ defined as

$$\zeta_t(\theta^n) \triangleq \begin{bmatrix} u_t \\ \varepsilon_t(\theta^n) \end{bmatrix}.$$

The importance of the latter, as will become clear in a moment, is to allow (by use of (5)) the prediction error gradient $\psi_t(\theta^n)$ to be expressed as

$$\begin{aligned} \psi_t(\theta^n) &= \frac{d\hat{y}_t(\theta^n)}{d\theta^n} = H^{-1}(q, \theta^n) \Pi'(q, \theta^n) \zeta_t(\theta^n), \\ \Pi'(q, \theta^n) &\triangleq \left[\frac{dG(q, \theta^n)}{d\theta^n}, \frac{dH(q, \theta^n)}{d\theta^n} \right]. \end{aligned} \quad (65)$$

Since this strategy of avoiding bias can lead to over-parameterisation of $G(q, \theta^n)$, the degree of which depends on how many extra terms in $A(q, \theta^n)$ are required for modeling $H(q)$, there is great importance in choosing the zeros $\{\xi_k\}$ not, in fact, all at the origin but as close as possible to where the true zeroes of $H(q)$ are believed to lie. For example, if $H(q)$ has a zero near the unit circle, an AR expansion $q^n/A(q, \theta^n)$ accounting for this zero will need to be of quite high order n before it is accurate [39].

A main purpose of this section is to highlight that unfortunately (and in resonance with the fixed denominator case) if according to this motivation the $\{\xi_k\}$ are not all chosen at the origin, then the approximation (64) can be quite inaccurate, even for large model order and data length.

This perhaps unexpected phenomenon can be illustrated in a fashion similar to that of § 2 by considering the least squares method (7) and an $n=8$ 'th order ARX-like structure (63) for the estimation of the ‘Åström system’

$$G(q) = \frac{q + 0.5}{q^2 - 1.5q + 0.7}, \quad H(q) = 1$$

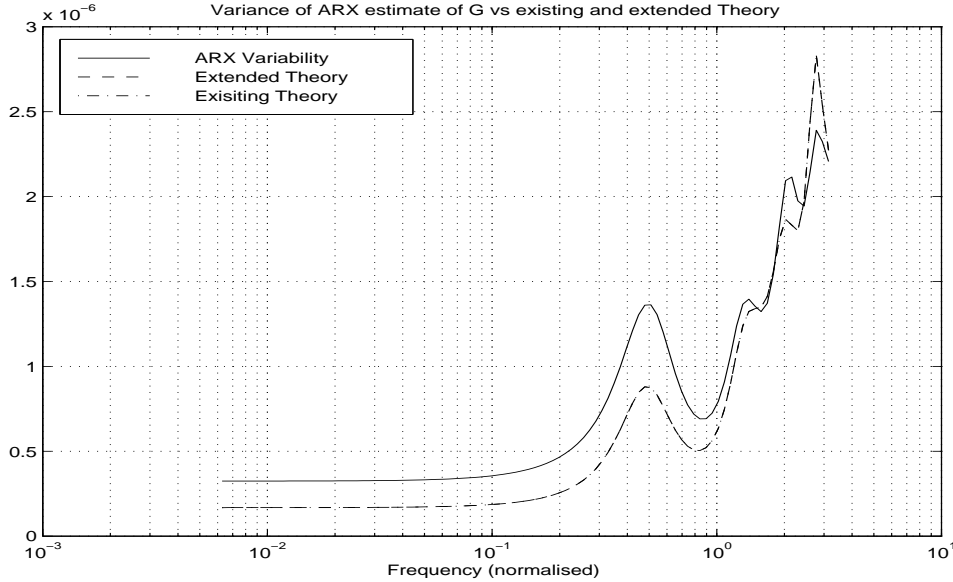


Figure 6: Conventional ARX with all noise model zeros at the origin. This is a comparison of Monte–Carlo estimate of sample variability (solid line) with (dash-dot line) the approximate expression (64). Note that this last line obscures a dashed line which is the new approximation (78) because for the case of all poles at the origin the pre-existing approximation (64) and the new one (78) are identical.

on the basis of $N = 10000$ observed open loop input-output measurements, the former being white Gaussian noise with spectral density $\Phi_u(\omega) = 0.25/(1.25 - \cos \omega)$ and the latter being corrupted by white Gaussian noise of variance $\sigma^2 = 0.001$. Suppose also, that all fixed noise model zeroes $\{\xi_k\}$ in $D_n(q)$ are chosen at the origin, so that a true ARX structure is employed. Note that in this example, and all the rest following in this section, the bias error in the estimation process is negligible, and hence the variance error will represent the total estimation error. In any event, since both $N = 10000$ and $n = 8$ can reasonably be considered large [22], then the approximation (64) for the variance error could be expected to be accurate, and indeed it appears to be so when shown as the dash-dot line in figure 6, with the sample average (over 500 Monte–Carlo simulations) estimate of the true variability being shown as the solid line.

However, if three noise model zeros are moved away from the origin to be at $\xi_k = \{0.8, 0.8, 0.8\}$, then the ensuing comparison of the theoretical (dash-dot line) approximation (64) and the Monte–Carlo estimate (solid line) of true variability shown in figure 7 shows much less agreement. Continuing, by choosing six noise model zeros away from the origin at $\{\xi_k\} = \{0.8, 0.8, 0.8, 0.7, 0.7, 0.7\}$, the results of this choice shown in figure 8 indicate that now the approximation between (dash-dot line) the theoretical approximation (64) and (solid line) the estimated true variability is so poor as to be considered very un-informative.

In contrast, the dashed line in figures 6–8 (which, in figure 6) is equal to and hence obscured by the dash-dot line) remains a good approximation regardless of the fixed zero position. This line is in fact the new approximation to be derived in this section, and again in resonance with the preceding fixed denominator analysis, it involves replacing the n term in (64) with a term $K_n(\omega, \omega)$. Since the latter in fact equals n for all the zeros at the origin, the original approximation (64) is encompassed as a special case.

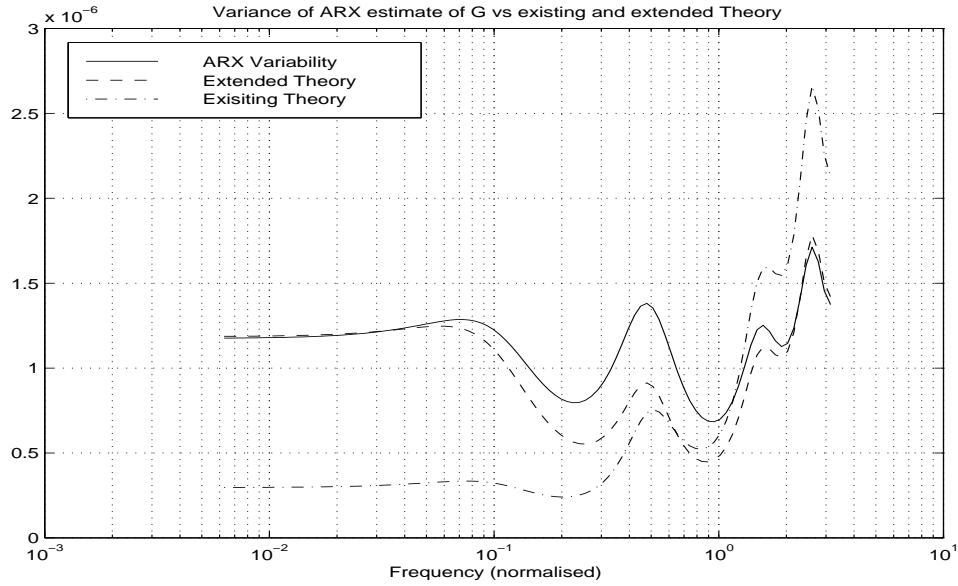


Figure 7: ARX-type structure with three noise model zeros not at the origin: comparison of Monte-Carlo estimate of sample variability (solid line) with (dash-dot line) the approximate expression (64) and (dashed line) the new approximation (79).

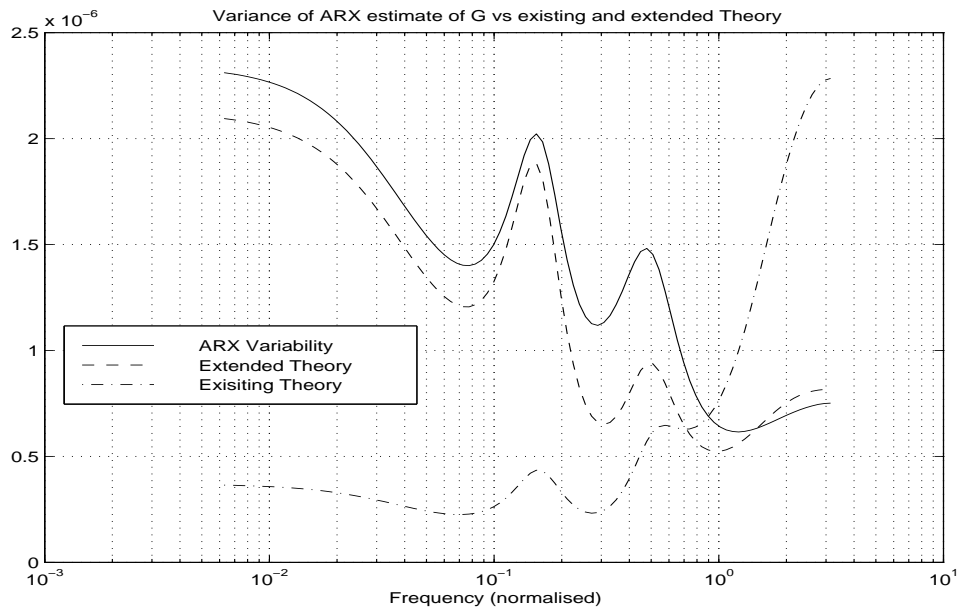


Figure 8: ARX-type structure with six noise model zeros not at the origin.

Also in accordance with the previous analysis, the improved approximation shown in figures 6–8 is obtained by re-parameterising the model structure into an equivalent orthonormal form which is tractable to analysis using the new generalised Fourier and Toeplitz results of [35] presented in § 5.

It is, of course, first necessary to establish the theoretical genesis of the inaccuracy phenomenon illustrated in figures 6–8, and for this purpose it is necessary to give a brief synopsis of the methods used in [22] to derive (64).

The conditions assumed there which are salient to the model structure are that it satisfy what is called a ‘shift’ property formulated as the requirement that with $\Gamma_n(q)$ defined as in (18), an $s \times s$ matrix of transfer functions $Z(q, \theta^n)$ exists such that

$$\Pi'(q, \theta^n) = [\Gamma_n(q) \otimes I_s] Z(q, \theta^n) \quad (66)$$

where the dimension s will be made clear in a moment by example and \otimes denotes the Kronecker tensor product of matrices defined for an $m \times n$ matrix A and an $\ell \times p$ matrix B to provide the $n\ell \times mp$ matrix $A \otimes B$ as

$$A \otimes B \triangleq \begin{pmatrix} a_{11}B & a_{12}B & \cdots & a_{1n}B \\ a_{21}B & a_{22}B & \cdots & a_{2n}B \\ \vdots & & & \vdots \\ a_{m1}B & a_{m2}B & \cdots & a_{mn}B \end{pmatrix}.$$

In any event, it is argued in [22] that assuming that θ^n is such that $G(q, \theta^n) = G(q)$, $H(q, \theta^n) = H(q)$ (that is, it is assumed that the true system is in the model set) then with the Toeplitz matrix definition (18),(20) extended to block-Toeplitz matrices via

$$T_n(F) \triangleq \frac{1}{2\pi} \int_{-\pi}^{\pi} [\Gamma_n(\omega) \otimes I_s] F(\omega) [\Gamma_n^*(\omega) \otimes I_s] d\omega \quad (67)$$

to handle $s \times s$ multi-variable spectral densities $F(\omega)$, then using known results on the asymptotic nature of block-Toeplitz matrices [49, 14]

$$P_n \sim T_n \left([Z(e^{j\omega}, \theta^n) \Phi_\zeta(\omega, \theta^n) Z^*(e^{j\omega}, \theta^n)]^{-1} \right) \text{ as } n \rightarrow \infty \quad (68)$$

so that recognising implicit Cesàro means and employing classical results from Fourier analysis via (14), it is argued in [22] that for large n the following approximations are valid (some abuse of notation is made in the interests of readability)

$$\begin{aligned} \frac{\Lambda_n(\omega)}{n} &= n^{-1} [\Pi'(e^{j\omega}, \theta^n)]^* P_n \Pi'(e^{j\omega}, \theta^n) \\ &\approx n^{-1} Z_o^*(\omega) [\Gamma_n^*(\omega) \otimes I_s] T_n \left([Z_o(\omega) \Phi_\zeta(\omega, \theta^n) Z_o^*(\omega)]^{-1} \right) [\Gamma_n(\omega) \otimes I_s] Z_o(\omega) \end{aligned} \quad (69)$$

$$\begin{aligned} &\approx |H_o(\omega)|^2 Z_o^*(\omega) [Z_o(\omega) \Phi_\zeta(\omega, \theta^n) Z_o^*(\omega)]^{-1} Z_o(\omega) \\ &= \sigma^2 |H(e^{j\omega}, \theta^n)|^2 \Phi_\zeta^{-1}(\omega, \theta^n). \end{aligned} \quad (70)$$

where in progressing to the last line, invertibility of $Z_o(\omega)$ has been assumed.

However, for the model structure (63), it is easily shown [22] that with the elements in the vector θ^n grouped into n blocks each of dimension $s = 2$ and each of the form $\theta_k^n = [a_k, b_k]^T$, then the corresponding $Z(q, \theta^n)$ of (66) is given as

$$Z(q, \theta^n) = \frac{1}{D_n(q)} \begin{bmatrix} -G(q, \theta^n) H(q, \theta^n) & -H^2(q, \theta^n) \\ H(q, \theta^n) & 0 \end{bmatrix}.$$

Therefore, when evaluated at θ^n it is better labeled as $Z_o(\omega, n)$ since if all the ξ_k are chosen away from the origin then via the $D_n^{-1}(q)$ term, the magnitude of the transfer function components of $Z(e^{j\omega}, \theta^n)$ all depend on n . In this case the Fourier convergence argument allowing the progression from (69) to (70) is problematic since the matrix valued function $[Z_o(\omega, n)\Phi_\zeta(\omega, \theta^n)Z_o^*(\omega, n)]^{-1}$ whose n 'th order Fourier reconstruction is implicit in (69), cannot be guaranteed to converge as n increases since the function possibly (depending on how many zeros of $D_n(q)$ are not at the origin) becomes less smooth with increasing n .

In the strict ARX case originally analysed in [22] there is (as illustrated in figure 6) no problem since in this instance, with all the zeros at the origin, $|D_n(e^{j\omega})| = |e^{j\omega n}| = 1$ so that by the assumption of the true system being in the model class, then the function $Z_o(\omega)\Phi_\zeta(\omega, \theta^n)Z_o^*(\omega)$ being implicitly Fourier reconstructed in (69) is invariant to increasing n , so that indeed it is valid to conclude increasing accuracy with increasing n and hence pass from (69) to (70).

Therefore, in order to provide the improved approximation shown in figures 6–8 for the case of many zeros of $D_n(q)$ not at the origin, the challenge is to mimic this invariant to n behaviour of the underlying Fourier reconstructed function.

As in § 5 this is achieved by using the observation that, asymptotically in n , the frequency response properties of estimates are invariant to (injective) changes in model structure, so that instead of the ARX-like structure (63) the equivalent, orthonormally parameterised, model structure

$$G(q, \theta^n) = \frac{B(q, \theta^n)}{A(q, \theta^n)} = \left[\sum_{k=0}^{n-1} b_k \mathcal{B}_k(q) \right] \left[1 - \sum_{k=0}^{n-1} a_k \mathcal{B}_k(q) \right]^{-1} \quad (71)$$

$$H(q, \theta^n) = \frac{1}{A(q, \theta^n)} = \left[1 - \sum_{k=0}^{n-1} a_k \mathcal{B}_k(q) \right] \quad (72)$$

is considered, whether or not the actual implementation employs this or the more natural form (63).

The advantage of the orthonormally parameterised form (71),(72) is that with the generalised definition (32) for $\Gamma_n(q)$, it satisfies a generalised shift property (again with $s = 2$)

$$\Pi'(q, \theta^n) = [\Gamma_n(q) \otimes I_s] Z(q, \theta^n)$$

where now

$$Z(q, \theta^n) = \begin{bmatrix} -G(q, \theta^n)H(q, \theta^n) & -H^2(q, \theta^n) \\ H(q, \theta^n) & 0 \end{bmatrix} \quad (73)$$

is (by the assumption of true system lying in the model class) *independent* of n for any choice of $\{\xi_k\}$.

To make this more precise, it is necessary to expand (31) to the definition of a generalised block-Toeplitz matrix depending on an $s \times s$ dimensional positive definite matrix valued function $F(\omega)$ as

$$M_n(F) \triangleq \frac{1}{2\pi} \int_{-\pi}^{\pi} [\Gamma_n(\omega) \otimes I_s] F(\omega) [\Gamma_n^*(\omega) \otimes I_s] d\omega \quad (74)$$

where $\Gamma_n(q)$ is the generalised form (32). Also, use the assumption that (at least asymptotically as $n \rightarrow \infty$) the true system is in the model set [22], then $\varepsilon_t(\theta^n) = e_t$ so that using (65), (10), (11) and Parseval's Theorem

$$R_n = M_n \left(\frac{Z(e^{j\omega}, \theta^n)\Phi_\zeta(\omega, \theta^n)Z^*(e^{j\omega}, \theta^n)}{|H(e^{j\omega}, \theta^n)|^2} \right), \quad Q_n = \sigma^2 R_n$$

and therefore

$$P_n = \sigma^2 R_n^{-1} = \sigma^2 M_n^{-1} \left(\frac{Z(e^{j\omega}, \theta_o^n) \Phi_\zeta(\omega, \theta_o^n) Z^*(e^{j\omega}, \theta_o^n)}{|H(e^{j\omega}, \theta_o^n)|^2} \right).$$

However, by Lemma A.5 the same equivalence (45) shown in [35] to hold for generalised Toeplitz matrices can also be shown to hold for generalised block Toeplitz matrices. In particular, using the definition (74)

$$M_n^{-1} \left(\frac{Z(e^{j\omega}, \theta_o^n) \Phi_\zeta(\omega, \theta_o^n) Z^*(e^{j\omega}, \theta_o^n)}{|H(e^{j\omega}, \theta_o^n)|^2} \right) \sim M_n (|H(e^{j\omega}, \theta_o^n)|^2 [Z(e^{j\omega}, \theta_o^n) \Phi_\zeta(\omega, \theta_o^n) Z^*(e^{j\omega}, \theta_o^n)]^{-1}) \quad (75)$$

as $n \rightarrow \infty$ where here the definition of the \sim equivalence is extended from that of (39), (40) to being that (75) implies that

$$\lim_{n \rightarrow \infty} K_n^{-1}(\omega, \omega) [\Gamma_n^*(\omega) \otimes I_s] [M_n(F^{-1}) - M_n^{-1}(F)] [M_n(F^{-1}) - M_n^{-1}(F)]^T [\Gamma_n(\omega) \otimes I_s] = 0$$

component-wise and for all $\omega \in [-\pi, \pi]$. Therefore, again using the reasoning in (41), then from (14) use of the orthonormal reparameterisation idea means that instead of (69) the following expression is considered

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} &= \lim_{n \rightarrow \infty} \frac{1}{K_n(\omega, \omega)} Z_o^*(\omega) [\Gamma_n^*(\omega) \otimes I_s] \times \\ &M_n \left(|H_o(\omega)|^2 [Z_o(\omega) \Phi_\zeta(\omega, \theta_o^n) Z_o^*(\omega)]^{-1} \right) [\Gamma_n(\omega) \otimes I_s] Z_o(\omega) + \\ &\frac{o(\|\hat{\theta}_N^n - \theta_o^n\|^2)}{K_n(\omega, \omega)}. \end{aligned} \quad (76)$$

Asymptotic in n analysis of this expression then provides what are the last main technical results of the paper.

Theorem 7.1. *With $\hat{\theta}_N^n$ calculated via (7) using the model structures (63), (71), (72) or any other equivalent structure with the same fixed noise model zeroes $\{\xi_k\}$ all chosen to lie within the unit disk \mathbf{D} , then in the limit as $N \rightarrow \infty$*

$$\sqrt{N} \begin{pmatrix} G(e^{j\omega}, \hat{\theta}_N^n) - G(e^{j\omega}, \theta_o^n) \\ H(e^{j\omega}, \hat{\theta}_N^n) - H(e^{j\omega}, \theta_o^n) \end{pmatrix} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Lambda_n(\omega))$$

where with $K_n(\omega, \omega) \equiv K_n(e^{j\omega}, e^{j\omega})$ given by (28) and under the assumption that eventually, for large enough n , the true system is in the model set, then

$$\lim_{n \rightarrow \infty} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} = \Phi_\nu(\omega) \begin{bmatrix} \Phi_u(\omega) & \Phi_{ue}(\omega) \\ \Phi_{ue}^*(\omega) & \sigma^2 \end{bmatrix}^{-1}$$

Proof. It has already been established in (13) that with the definition of the composite transfer matrix (12)

$$\sqrt{N} \left[\Pi(e^{j\omega}, \hat{\theta}_N^n) - \Pi(e^{j\omega}, \hat{\theta}_o^n) \right] \xrightarrow{\mathcal{D}} \mathcal{N}(0, \Lambda_n(\omega)) \text{ as } N \rightarrow \infty$$

where (76) gives an asymptotic in n expression for $\Lambda_n(\omega)/K_n(\omega, \omega)$. Because of the Kronecker product terms in (76), the result (50) is not applicable. However, Lemma A.3 shows that for $F(\omega)$ an arbitrary regular multi-variable spectral density (50) may be extended to

$$\lim_{n \rightarrow \infty} \frac{1}{K_n(\omega, \omega)} [\Gamma_n^*(\omega) \otimes I_s] M_n(F) [\Gamma_n(\omega) \otimes I_s] = F(\omega)$$

component-wise in $F(\omega)$ and where $M_n(F)$ is defined by (67). Applying this result to (76) with $F(\omega) = |H_o(\omega)|^2 [Z_o(\omega)\Phi_\zeta(\omega, \theta_o^n)Z_o^*(\omega)]^{-1}$ then gives (since $Z(e^{j\omega}, \theta_o^n)$ is invertible)

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{\Lambda_n(\omega)}{K_n(\omega, \omega)} &= |H_o(\omega)|^2 Z_o^*(\omega) [Z_o(\omega)\Phi_\zeta(\omega, \theta_o^n)Z_o^*(\omega)]^{-1} Z_o(\omega) \\ &= \sigma^2 |H(e^{j\omega}, \theta_o^n)|^2 \Phi_\zeta^{-1}(\omega, \theta_o^n). \end{aligned}$$

□

Corollary 7.1. *Under the same conditions as the previous Theorem, but with a strengthened requirement that $\mathbf{E}\{e_k^8\} < \infty$ then*

$$\lim_{n \rightarrow \infty} \lim_{N \rightarrow \infty} \frac{N}{K_n(\omega, \omega)} \text{Cov} \left\{ \Pi(e^{j\omega}, \hat{\theta}_N^n) \right\} = \Phi_\nu(\omega) \begin{bmatrix} \Phi_u(\omega) & \Phi_{ue}(\omega) \\ \Phi_{ue}^*(\omega) & \sigma^2 \end{bmatrix}^{-1} \quad (77)$$

Proof. Follows from the previous Theorem using the methods in Appendix 9B of [20]. □

The essential point of Theorem 7.1 and Corollary 7.1 that discriminates them from previous corresponding results [22] is that the asymptotic expression (52) implicitly involves generalised Fourier reconstruction of a matrix valued function $|H_o(\omega)|^2 [Z_o(\omega)\Phi_\zeta(\omega, \theta_o^n)Z_o^*(\omega)]^{-1}$ which is *invariant* to n , regardless of whether fixed noise model zeros are chosen at the origin or not. In this case, and unlike the analysis leading to (71), it can be expected that the approximation between the left hand side of (52) and its asymptotic value on the right hand side of (52) will monotonically improve with increasing n , again regardless of the location of the fixed noise model zeros.

As a consequence of this the suggestion of this paper is that an improved approximation for the variability of ARX-type model structure estimates is obtained from (52) by assuming that for finite N and n convergence has approximately converged in (52) so that

$$\text{Cov} \left\{ \Pi(e^{j\omega}, \hat{\theta}_N^n) \right\} \approx \frac{1}{N} \Phi_\nu(\omega) \begin{bmatrix} \Phi_u(\omega) & \Phi_{ue}(\omega) \\ \Phi_{ue}(\omega) & \sigma^2 \end{bmatrix}^{-1} \sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2. \quad (78)$$

In particular, if the data is collected in open loop ($\Phi_{ue}(\omega) = 0$) then

$$\text{Cov} \left\{ G(e^{j\omega}, \hat{\theta}_N^n) \right\} \approx \frac{1}{N} \frac{\Phi_\nu(\omega)}{\Phi_u(\omega)} \sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2 \quad (79)$$

and

$$\text{Cov} \left\{ H(e^{j\omega}, \hat{\theta}_N^n) \right\} \approx \frac{|H(e^{j\omega}, \theta_o^n)|^2}{N} \sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2 \quad (80)$$

which explicitly shows, via the $\sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2 = K_n(\omega, \omega)$ term how the choice of fixed zeros $\{\xi_k\}$ in the noise model influences the sensitivity of the final transfer function estimation to measurement noise; again, see figure 4. Note that as per the previous section on fixed denominator model structures, these new expressions contain the results for equation error structures presented in [22] as a special case of $\xi_k = 0$ since the latter implies $\sum_{k=0}^{n-1} |\mathcal{B}_k(e^{j\omega})|^2 = n$ so that (78) becomes (64).

To conclude this section, it is interesting to note two important cases where fixed zero noise models in ARX structures are implicitly involved, albeit via alternate motivations. Firstly, consider the case of employing a standard ARX structure save that, in the interests of concentrating attention

on certain frequency regions [46], the common practice of data pre-filtering is employed in such a way that (denoting the filtering action by the transfer function $F(q)$) the ARX model structure being employed becomes

$$A(q, \theta^n)F(q)y_t = B(q, \theta^n)F(q)u_t + e_t$$

In this case, if the filter $F(q)$ is all pole of the form $F(q) = 1/D_n(q)$ then this model structure is clearly identical to the model structure (63) with fixed-zero noise model and for which it was shown in figures 7–8 that the approximation (64) can be quite poor if the number of pre-filter poles is an appreciable proportion of the model order. This leads to the conclusion that whenever all-pole pre-filters are used with ARX structures, then the improved approximation (78) is preferable.

Finally, in the interests of numerical conditioning when sampling well beyond the Nyquist rate, the so-called ‘ δ operator’ model structure is an option [30, 7]. Such structures, when in equation error form as presented in [30] appear as

$$\frac{A(\delta, \theta^n)}{D_n(\delta)}y_t = \frac{B(\delta, \theta^n)}{D_n(\delta)}u_t + e_t \quad (81)$$

where with Δ equal to the sampling period in seconds, δ is defined as

$$\delta \triangleq \frac{q - 1}{\Delta}$$

and

$$\begin{aligned} A(\delta) &= a_0 + a_1\delta + a_2\delta^2 + \cdots + a_{n-1}\delta^{n-1} + \delta^n, \\ B(q) &= b_0 + b_1\delta + b_2\delta^2 + \cdots + b_{n-1}\delta^{n-1} \end{aligned}$$

with

$$D_n(q) = \prod_{k=1}^n (\delta - \tilde{\xi}_k)$$

and now the ξ'_k are chosen to all lie in the disc $\tilde{\mathbf{D}} \triangleq \{z \in \mathbf{C} : |z + 1/\Delta| < 1/\Delta\}$. Aside from numerical considerations, another dividend of using this model structure is that since the δ operator is the Euler differencing approximation to the differentiation operator, the ensuing estimates can be interpreted as an approximation of the underlying continuous time system. Again, this depends on sampling well beyond the Nyquist rate [32].

In the q operator case and in the absence of noise model knowledge it is natural to simply choose $D_n(q) = q^n$. However, when operating in a δ operator framework it is not natural to make the specific equivalent choice $D_n = (\delta + 1/\Delta)^n$, since a key philosophy underlying the use of δ operators is that of more closely approaching continuous time intuition wherein the precise value of Δ is incidental. Other considerations therefore come into play in the choice of the zeros of $D_n(\delta)$, for which [30] provides a discussion.

In this setting, attention is therefore focused on a model structure (81) which is (modulo numerical issues) equivalent to the shift operator model (63) with fixed noise model zeros not at the origin, but at the locations $\xi_k = \Delta\tilde{\xi}_k + 1$. The most accurate quantification of the variability of the ensuing δ operator based estimates will therefore be given by the new approximation (78). These δ operator inspired asymptotic variability issues were first prescinded in [10].

8 Conclusion

This paper has provided an exposition of the previously unsuspected intrinsic nature of orthonormal bases in the study of certain least-squares estimation problems. In works [25, 22, 20] preceding this paper, the bases have been present in the special form corresponding to $\xi_k = 0$ of $\mathcal{B}_n(e^{j\omega}) = e^{-j\omega n}$ so that they have been hidden since then $|\mathcal{B}_n(e^{j\omega})|^2 = |e^{-j\omega n}|^2 = 1$. However, as shown for the first time here in a completely general setting, they become apparent for certain structures with fixed pole or zero terms where $|D_n(e^{j\omega})|^2 \neq 1$ and hence whose changing nature with respect to increasing n cannot be ignored in arguments that are asymptotic in n . This paper has shown that the analytic key to circumventing this difficulty is to develop and apply results that generalise certain Fourier convergence and asymptotic Toeplitz matrix properties to the case of an underlying general rational orthonormal basis, of which the classical trigonometric basis $\{e^{-j\omega n}\}$ is a special case.

As a final comment, it is important to acknowledge that the orthonormal basis ideas used in this paper have a very long history. The study of the formulation (25) in mathematical literature can be traced back at least as far as [28, 42], and the engineering applications (25) seem to originate in [17], although it clear that Wiener was aware of the engineering relevance of (25) somewhat earlier [19]. In more recent times, the list of works employing specialised ‘Laguerre’ or ‘two-parameter Kautz’ forms of (25) is immense, but a partial list for readers interested in further investigations should include [29, 4, 38, 36, 26, 44, 45, 37, 15, 6, 36, 18]. However, a key feature discriminating this work from that in the current paper, is that in pre-existing work the orthonormal bases were explored as an *implementational* option while here, they have been employed purely as an *analysis* tool, that is applicable to model structures that are not ab-initio formulated in terms of them.

A Auxiliary Results

Lemma A.1. *Let $\{\nu_t\}$ be a stationary stochastic process with associated covariance function $R_\nu(\tau)$ and spectral density $\Phi_\nu(\omega)$ satisfying*

$$\Phi_\nu(\omega) = \sum_{\tau=-\infty}^{\infty} R_\nu(\tau) e^{-j\omega\tau}, \quad \sum_{\tau=-\infty}^{\infty} |R_\nu(\tau)| < \infty,$$

and suppose $\{u_t\}$ is a quasi-stationary process with associated spectral density $\Phi_u(\omega)$. Then with $\psi_t = \Gamma_n(q)u_t$ and $\Gamma_n(q)$ defined by (32)

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T \} \mathbf{E} \{ \nu_k \nu_m \} = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma(\omega) \Gamma^*(\omega) \Phi_u(\omega) \Phi_\nu(\omega) d\omega.$$

Proof. Without loss of generality assume $u_t = 0$ for $t < 0$ so that

$$\begin{aligned} \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T \} \mathbf{E} \{ \nu_k \nu_m \} &= \lim_{N \rightarrow \infty} \frac{1}{2\pi N} \int_{-\pi}^{\pi} \Phi_\nu(\omega) \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \mathbf{E} \{ \psi_k \psi_m^T \} e^{j\omega(k-m)} d\omega \\ &= \lim_{N \rightarrow \infty} \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_\nu(\omega) \sum_{\tau=1-N}^{N-1} \underbrace{\left(\frac{1}{N} \sum_{k=0}^{N-1} \mathbf{E} \{ \psi_k \psi_{k-|\tau|}^T \} \right)}_{R_N(\tau)} e^{j\omega\tau} d\omega. \end{aligned}$$

Therefore, defining

$$R_\psi(\tau) = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{k=0}^{N-1} \mathbf{E} \left\{ \psi_k \psi_{k-|\tau|}^T \right\} = \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma_n(\omega) \Gamma_n^*(\omega) \Phi_u(\omega) e^{-j\omega\tau} d\omega$$

leads to (the absolute value inequalities following are to be interpreted component-wise in the matrix quantities involved)

$$\begin{aligned} & \left| \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma(\omega) \Gamma^*(\omega) \Phi_u(\omega) \Phi_\nu(\omega) d\omega - \frac{1}{N} \sum_{k=0}^{N-1} \sum_{m=0}^{N-1} \psi_k \psi_m^T R_\nu(k-m) \right| = \\ & \left| \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_\nu(\omega) \left[\left(\sum_{\tau=1-N}^{N-1} R_\psi(\tau) - R_N(\tau) \right) + \sum_{|\tau| \geq N} R_\psi(\tau) \right] e^{j\omega\tau} d\omega \right| \\ & \leq \sum_{\tau=1-N}^{N-1} |R_\psi(\tau) - R_N(\tau)| |R_\nu(\tau)| + \left| \sum_{|\tau| \geq N} R_\psi(\tau) R_\nu(\tau) \right| \\ & \leq \left(\max_{|\tau| < N} |R_\psi(\tau) - R_N(\tau)| \right) \sum_{\tau=1-N}^{N-1} |R_\nu(\tau)| + \left(\sup_{|\tau| \geq N} |R_\nu(\tau)| \right) \sum_{|\tau| \geq N} |R_\psi(\tau)|. \end{aligned}$$

However, by the quasi-stationarity assumption on $\{u_t\}$, then $\lim_{N \rightarrow \infty} R_N(\tau) = R_\psi(\tau)$ so for sufficiently large N then $\max_{|\tau| < N} |R_\psi(\tau) - R_N(\tau)|$ can be made arbitrarily small, so since also by assumption $\sum_{\tau} |R_\nu(\tau)| < \infty$, then the first term in the above overbound tends to zero with increasing N . For the second term, note that since by the assumption that $\sum_{\tau} |R_\nu(\tau)| < \infty$, then $\limsup_{\tau \rightarrow \infty} |R_\nu(\tau)| = 0$ so that for large enough N , then $\sup_{|\tau| \geq N} |R_\nu(\tau)|$ can be made arbitrarily small. As well, by the quasi-stationarity assumption on $\{u_t\}$, which implies that $\sum_{\tau=-\infty}^{\infty} R_\psi(\tau)$ exists component-wise, then $\limsup_{N \rightarrow \infty} \sum_{|\tau| \geq N} |R_\psi(\tau)| = 0$, which completes the proof. \square

Lemma A.2. *Suppose that $\{u_t\}$ and $\{e_t\}$ are both realisations of zero mean Gaussian distributed stationary stochastic processes and the orthonormal model structure (27) is employed. Then regardless of whether the data is collected in open or closed loop, and with no approximation caused by neglecting an undermodelling induced term, the matrix Q_n defined in (11) obeys*

$$Q_n \sim M_n(\Phi_u(\omega) \Phi_\varepsilon(\omega, \theta_\circ^n))$$

as $n \rightarrow \infty$ where the matrix formulation M_n is defined in (31) and $\Phi_\varepsilon(\omega, \theta_\circ^n)$ is the spectral density of the prediction residuals $\{\varepsilon_t(\theta_\circ^n)\}$ defined in (6) and evaluated at $\theta^n = \theta_\circ^n$ defined in (31).

Proof. From the definition (11) after using the stationarity assumption and the change of variable $\tau = k + m$ and re-grouping terms

$$\begin{aligned} Q_n &= \lim_{N \rightarrow \infty} \sum_{\tau=-N}^N \left(1 - \frac{|\tau|}{N} \right) \mathbf{E} \left\{ \psi_k \psi_{k-\tau}^T \varepsilon_k(\theta_\circ^n) \varepsilon_{k-\tau}(\theta_\circ^n) \right\}, \\ &= \sum_{\tau=-\infty}^{\infty} \mathbf{E} \left\{ \psi_k \psi_{k-\tau}^T \varepsilon_k(\theta_\circ^n) \varepsilon_{k-\tau}(\theta_\circ^n) \right\}, \end{aligned}$$

where the properties of Cesàro means have been used in progressing to the last line. Using the Gaussianity assumption and the formula for fourth moments of jointly Gaussian random variables [43]

$$\mathbf{E} \{ \psi_k \psi_{k-\tau}^T \varepsilon_k(\theta_\circ^n) \varepsilon_{k-\tau}(\theta_\circ^n) \} = \mathbf{E} \{ \psi_k \psi_{k-\tau}^T \} \mathbf{E} \{ \varepsilon_k(\theta_\circ^n) \varepsilon_{k-\tau}(\theta_\circ^n) \} + \mathbf{E} \{ \psi_k \varepsilon_{k-\tau}(\theta_\circ^n) \} \mathbf{E} \{ \psi_{k-\tau}^T \varepsilon_k(\theta_\circ^n) \}$$

where use is made of the fact that by the definition of θ_\circ^n , $\mathbf{E} \{ \psi_k \varepsilon_k(\theta_\circ^n) \} = 0$. Furthermore, (suppressing the dependence on θ_\circ^n and with $\delta(\cdot)$ denoting Kronecker delta)

$$\begin{aligned} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \psi_{k-\tau}^T \} \mathbf{E} \{ \varepsilon_k \varepsilon_{k-\tau} \} &= \sum_{\ell=-\infty}^{\infty} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \psi_{k-\tau}^T \} \mathbf{E} \{ \varepsilon_k \varepsilon_{k-\ell} \} \delta(\ell - \tau), \\ &= \sum_{\ell=-\infty}^{\infty} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \psi_{k-\tau}^T \} \mathbf{E} \{ \varepsilon_k \varepsilon_{k-\ell} \} \frac{1}{2\pi} \int_{-\pi}^{\pi} e^{j\omega(\ell-\tau)} d\omega, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \psi_{k-\tau}^T \} e^{-j\omega\tau} \sum_{\ell=-\infty}^{\infty} \mathbf{E} \{ \varepsilon_k \varepsilon_{k-\ell} \} e^{j\omega\ell} d\omega, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_\psi(\omega) \overline{\Phi_\varepsilon(\omega)} d\omega, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma_n(\omega) \Gamma_n^*(\omega) \Phi_u(\omega) \Phi_\varepsilon(\omega) d\omega, \\ &= M_n(\Phi_u \Phi_\varepsilon). \end{aligned}$$

Using an identical line of argument

$$\begin{aligned} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \varepsilon_{k-\tau} \} \mathbf{E} \{ \psi_{k-\tau}^T \varepsilon_k \} &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \varepsilon_{k-\tau} \} e^{-j\omega\tau} \sum_{\ell=-\infty}^{\infty} \mathbf{E} \{ \psi_{k-\ell}^T \varepsilon_k \} e^{j\omega\ell} d\omega, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \sum_{\tau=-\infty}^{\infty} \mathbf{E} \{ \psi_k \varepsilon_{k-\tau} \} e^{-j\omega\tau} \sum_{\ell=-\infty}^{\infty} \mathbf{E} \{ \psi_{k-\ell}^T \varepsilon_{k-\ell} \} e^{-j\omega\ell} d\omega, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_{\psi_\varepsilon}(\omega) \Phi_{\psi_\varepsilon}^T(\omega) d\omega, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \Gamma_n(\omega) \Gamma_n^T(\omega) \Phi_{u\varepsilon}^2(\omega) d\omega \triangleq X_n. \end{aligned}$$

Now, in the case of all the poles $\{\xi_k\}$ being real, then since $\mathcal{B}_k(e^{j\omega}) = \overline{\mathcal{B}_k(e^{-j\omega})}$

$$\left| \frac{\Gamma_n^*(\omega) X_n \Gamma_n(\omega)}{K_n(\omega, \omega)} \right| \leq \frac{\|\Phi_{u\varepsilon}\|_\infty}{2\pi K_n(\omega, \omega)} \int_{-\pi}^{\pi} |K_n(\lambda, \omega)| |K_n(-\omega, \lambda)| d\lambda.$$

However, in [35] it is shown that expressions of this form tend to zero with increasing n provided all the poles $\{\xi_k\}$ are chosen within the open disc \mathbf{D} . This same result can also be shown to hold for the $\{\xi_k\}$ being complex, but at the expense of considerably more involved arithmetic which is not appropriate to document here. Therefore, since $Q_n = M_n(\Phi_u \Phi_\varepsilon) + X_n$ and it has just been established that $X_n \sim 0$ as $n \rightarrow \infty$, then $Q_n \sim M_n(\Phi_u \Phi_\varepsilon)$ as $n \rightarrow \infty$. \square

Lemma A.3. Let $F(\omega)$ be a continuous $s \times s$ matrix valued function defined on $[-\pi, \pi]$. Then

$$\lim_{n \rightarrow \infty} \frac{1}{K_n(\omega, \omega)} [\Gamma_n^*(\omega) \otimes I_s] M_n(F) [\Gamma_n(\omega) \otimes I_s] = F(\omega)$$

component-wise in $F(\omega)$ and where $M_n(F)$ is defined by (67) with $\Gamma_n(\omega)$ being the generalised form (32).

Proof. Using the algebraic properties of Kronecker tensor product [3] and the definition (67) of $M_n(F)$

$$\begin{aligned} [\Gamma_n^*(\omega) \otimes I_s] M_n(F) [\Gamma_n(\omega) \otimes I_s] &= [\Gamma_n^*(\omega) \otimes I_s] \left[\frac{1}{2\pi} \int_{-\pi}^{\pi} [\Gamma_n(\lambda) \otimes I_s] F(\lambda) [\Gamma_n^*(\lambda) \otimes I_s] d\lambda \right] \times \\ &\quad [\Gamma_n(\omega) \otimes I_s], \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} [\Gamma_n^*(\omega) \otimes I_s] [\Gamma_n(\lambda) \Gamma_n^*(\lambda) \otimes F(\lambda)] [\Gamma_n(\omega) \otimes I_s] d\lambda, \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} |\Gamma_n^*(\omega) \Gamma_n(\lambda)|^2 \otimes F(\lambda) d\lambda. \end{aligned}$$

Use of the result (50) on noting that the above expression is a matrix made up of terms of the form $\Gamma_n^*(\omega) M_n(F_{mn}) \Gamma_n(\omega)$ with $F_{mn}(\lambda)$ being the m, n 'th scalar entry of $F(\lambda)$ and with $M_n(f)$ defined by (31),(32) then completes the proof. \square

Lemma A.4. *With the definition of generalised Toeplitz matrix being expanded to that of generalised block-Toeplitz matrix by (74), then for F, G any $s \times s$ complex matrix valued functions which have entries which are Lipschitz continuous of order $\varepsilon > 0$*

$$\lim_{n \rightarrow \infty} K_n^{-1}(\omega, \omega) [\Gamma_n^*(\omega) \otimes I_s] [M_n(F) M_n(G) - M_n(FG)]^* [M_n(F) M_n(G) - M_n(FG)] [\Gamma_n(\omega) \otimes I_s] = 0 \quad (\text{A.1})$$

component-wise.

Proof. Using the formulation (74)

$$\begin{aligned} [M_n(F) M_n(G) [\Gamma_n(\omega) \otimes I_s]]_{rs+k,t} &= \\ &= \frac{1}{4\pi^2} \int_{-\pi}^{\pi} \int_{-\pi}^{\pi} [\Gamma_n(\lambda) K_n(\omega, \sigma) K_n(\sigma, \lambda) \otimes F(\lambda) G(\sigma)]_{rs+k,t} d\lambda d\sigma \\ &= \frac{1}{4\pi^2} \int_{-\pi}^{\pi} \int_{-\pi}^{\pi} \mathcal{B}_r(\lambda) K_n(\omega, \sigma) K_n(\sigma, \lambda) [F(\lambda) G(\sigma)]_{k,t} d\lambda d\sigma \\ &= \frac{1}{4\pi^2} \sum_{\ell=0}^{s-1} \int_{-\pi}^{\pi} \int_{-\pi}^{\pi} \mathcal{B}_r(\lambda) K_n(\omega, \sigma) K_n(\sigma, \lambda) [F(\lambda)]_{k,\ell} [G(\sigma)]_{\ell,t} d\lambda d\sigma \\ &= \sum_{\ell=0}^{s-1} [M_n(F_{k,\ell}) M_n(G_{\ell,t}) \Gamma_n(\omega)]_r. \end{aligned}$$

Also, by a similar argument

$$[M_n(FG) [\Gamma_n(\omega) \otimes I_s]]_{rs+k,t} = \sum_{\ell=0}^{s-1} [M_n(F_{k,\ell} G_{\ell,t}) \Gamma_n(\omega)]_r.$$

Therefore, defining $\Delta_n \triangleq M_n(F) M_n(G) - M_n(FG)$ leads to

$$[[\Gamma_n^*(\omega) \otimes I_s] \Delta_n^* \Delta_n [\Gamma_n(\omega) \otimes I_s]]_{p,t} =$$

$$\begin{aligned}
 & \sum_{r=0}^{n-1} \sum_{k=0}^{s-1} \overline{[\Delta_n[\Phi_n(\omega) \otimes I_s]]_{rs+k,p} [\Delta_n[\Phi_n(\omega) \otimes I_s]]_{rs+k,t}} \\
 = & \sum_{r=0}^{n-1} \sum_{k=0}^{s-1} \sum_{\ell=0}^{s-1} \sum_{v=0}^{s-1} \overline{[[M_n(F_{k,\ell})M_n(G_{\ell,p}) - M_n(F_{k,\ell}, G_{\ell,p})]\Gamma_n(\omega)]_r} \times \\
 & [[M_n(F_{k,\ell})M_n(G_{\ell,t}) - M_n(F_{k,\ell}, G_{\ell,t})]\Gamma_n(\omega)]_r
 \end{aligned}$$

Now, when $p = t$ then the scalar result (44) gives that this expression divided by $K_n(\omega, \omega)$ tends to zero as $n \rightarrow \infty$. When $p \neq t$ it is first necessary to upper-bound the above expression using the Cauchy–Schwarz inequality as $|\Gamma_n^* A B \Gamma_n|^2 \leq |\Gamma_n^* A A^* \Gamma_n| |\Gamma_n^* B^* B \Gamma_n|$ before again using the scalar result (44) to conclude that this overbound, and hence the expression of interest, also tends to zero when divided by $K_n(\omega, \omega)$ as $n \rightarrow \infty$. \square

Lemma A.5. *With the property (A.1) being understood as the defining feature of the relationship $M_n(F) \sim M_n(G)$ as $n \rightarrow \infty$ for generalised block-Toeplitz matrices defined by (74), then for $F(\omega)$ any Lipschitz continuous matrix valued function that is invertible*

$$M_n^{-1}(F) \sim M_n(F^{-1}) \text{ as } n \rightarrow \infty.$$

Proof.

$$\begin{aligned}
 & [\Gamma_n^*(\omega) \otimes I_s][M_n^{-1}(F) - M_n(F^{-1})]^T [M_n^{-1}(F) - M_n(F^{-1})][\Gamma_n(\omega) \otimes I_s] = \\
 & [\Gamma_n^*(\omega) \otimes I_s][I - M_n(F)M_n(F^{-1})]^T M_n^{-T}(F)M_n^{-1}(F)[I - M_n(F)M_n(F^{-1})][\Gamma_n(\omega) \otimes I_s]
 \end{aligned}$$

However, by lemma A.4 $M_n(F)M_n(F^{-1}) \sim M_n(I) = I$ as $n \rightarrow \infty$ so that use of the Cauchy–Schwarz inequality and the bound $\|M_n(F^{-1})\| \leq \|F^{-1}\|_\infty$ completes the proof. \square

References

- [1] *Special issue on system identification for robust control design*, IEEE Transactions on Automatic Control, 37 (1992).
- [2] R. BITMEAD, M. GEVERS, AND V. WERTZ, *Adaptive Optimal Control, The Thinking Man's GPC.*, Prentice Hall, Englewood Cliffs, New Jersey, 1990.
- [3] J. W. BREWER, *Kronecker products and matrix calculus in system theory*, IEEE Transactions on Circuits and Systems, 25 (1978), pp. 772–781.
- [4] P. W. BROOME, *Discrete orthonormal sequences*, Journal of the Association for Computing Machinery, 12 (1965), pp. 151–168.
- [5] P. CAINES, *Linear Stochastic Systems*, John Wiley and Sons, New York, 1988.
- [6] P. R. CLEMENT, *Laguerre functions in signal analysis and parameter identification*, Journal of the Franklin Institute, 313 (1982), pp. 85–95.
- [7] A. FEUER AND R. MIDDLETON, *Conditioning of LMS algorithms with fast sampling*, IEEE Transactions on Signal Processing, 43 (1995), pp. 1978–1981.

- [8] M. GEVERS, L. LJUNG, AND P.M.J. VAN DEN HOF, *Asymptotic variance expressions for closed-loop identification and their relevance in identification for control*, in Proceedings of the 11th IFAC Symposium on System Identification, 1997, pp. 1449–1454.
- [9] G. GOLUB AND C. V. LOAN, *Matrix Computations*, Johns Hopkins University Press, 1989.
- [10] G. GOODWIN, M. GEVERS, AND D. MAYNE, *Bias and variance distribution in transfer function estimates*, in Proceedings of 9th IFAC Symposium on Identification and System Parameter Estimation, Budapest, 1991, pp. 952–957.
- [11] U. GRENANDER AND G. SZEGÖ, *Toeplitz Forms and their Applications*, University of California Press, Berkeley, 1958.
- [12] L. GUO AND L. LJUNG, *The role of model validation for assessing the size of unmodelled dynamics*, in Proceedings of the 33rd Conference on Decision and Control, IEEE, December 1994, pp. 3894–3899.
- [13] E. HANNAN AND D. NICHOLLS, *The estimation of the prediction error variance*, Journal of the American Statistical Association, 72 (1977), pp. 834–840.
- [14] E. HANNAN AND B. WAHLBERG, *Convergence rates for inverse Toeplitz matrix forms*, Journal of Multivariate Analysis, 31 (1989), pp. 127–135.
- [15] P. HEUBERGER, P.M.J. VAN DEN HOF, AND O. BOSGRA, *A generalized orthonormal basis for linear dynamical systems*, IEEE Transactions on Automatic Control, AC-40 (1995), pp. 451–465.
- [16] H. HJALMARSSON, M. GEVERS, AND F. DE BRUYNE, *For some model based control design criteria, closed loop identification gives better performance*, Automatica, 32 (1996), pp. 1659–1673.
- [17] W. H. KAUTZ, *Network synthesis for specified transient response*, Tech. Rep. 209, Massachusetts Institute of Technology, Research Laboratory of Electronics, 1952.
- [18] ———, *Transient synthesis in the time domain*, IRE Transactions on Circuit Theory, 1 (1954), pp. 29–39.
- [19] Y. W. LEE, *Statistical Theory of Communication*, Wiley, New York, 1960.
- [20] L. LJUNG, *System Identification: Theory for the User*, Prentice-Hall, Inc., New Jersey, 1987.
- [21] L. LJUNG, *Identification in closed loop: Some aspects on direct and indirect approaches*, in Proceedings of the 11th IFAC Symposium on System Identification, 1997, pp. 141–146.
- [22] L.LJUNG, *Asymptotic variance expressions for identified black-box transfer function models*, IEEE Transactions on Automatic Control, AC-30 (1985), pp. 834–844.
- [23] L.LJUNG AND P.E.CAINES, *Asymptotic normality of prediction error estimators for approximate system models*, Stochastics, 3 (1979), pp. 29–46.
- [24] L.LJUNG AND B. WAHLBERG, *Asymptotic properties of the least squares method for estimating transfer functions and disturbance spectra*, Advances in Applied Probability, 24 (1992), pp. 412–440.

- [25] L.LJUNG AND Z.D.YUAN, *Asymptotic properties of black-box identification of transfer functions*, IEEE Transactions on Automatic Control, 30 (1985), pp. 514–530.
- [26] P. MÄKILÄ, *Laguerre methods and h_∞ identification of continuous-time systems*, International Journal of Control, 58 (1993), pp. 665–683.
- [27] P. MÄKILÄ, J. PARTINGTON, AND T. GUSTAFSSON, *Worst-case control-relevant identification*, Automatica, 31 (1995), pp. 1799–1820.
- [28] F. MALMQUIST, *Sur la détermination d'une classe de fonctions analytiques par leurs valeurs dans un ensemble donné de points*, Comptes Rendus du Sixième Congrès des mathématiciens scandinaves (Kopenhagen), (1925), pp. 253–259.
- [29] J. MENDEL, *A unified approach to the synthesis of orthonormal exponential functions useful in systems analysis*, IEEE Transactions on Systems Science and Cybernetics, 2 (1966), pp. 54–62.
- [30] R. MIDDLETON AND G. GOODWIN, *Digital Estimation and Control: A Unified Approach*, Prentice-Hall, Inc., New Jersey, 1990.
- [31] B. NINNESS, *Integral constraints on the accuracy of least squares estimation*, Automatica, 32 (1996), pp. 391–397.
- [32] B. NINNESS AND G. GOODWIN, *Identification of Continuous Time Systems*, Kluwer Academic Publications, New York., 1991, ch. The Relationship Between Discrete Time and Continuous Time Linear Estimation.
- [33] B. NINNESS AND G. GOODWIN, *Estimation of model quality*, Automatica, 31 (1995), pp. 32–74.
- [34] B. NINNESS AND F. GUSTAFSSON, *A unifying construction of orthonormal bases for system identification*, IEEE Transactions on Automatic Control, 42 (1997), pp. 515–521.
- [35] B. NINNESS, H. HJALMARSSON, AND F. GUSTAFSSON, *Generalised Fourier and Toeplitz results for rational orthonormal bases*, Technical Report EE9740, Department of Electrical and Computer Engineering, University of Newcastle, Australia. To appear in SIAM Journal on Control and Optimization, (1997).
- [36] Ü. NURGES, *Laguerre models in problems of approximation and identification*, Adaptive Systems, (1987), pp. 346–352. Published by Plenum Publishing, Translated from Avtomatika i Telemekhanika (3) pp 88-96, March 1987.
- [37] P.M.J. VAN DEN HOF, P.S.C. HEUBERGER, AND J. BOKOR, *System identification with generalized orthonormal basis functions*, Automatica, 31 (1995), pp. 1821–1834.
- [38] D. ROSS, *Orthonormal exponentials*, IEEE Transactions on Communication and Electronics, 71 (1964), pp. 173–176.
- [39] M. SALGADO, B. NINNESS, AND G. GOODWIN, *Generalised expansion algorithm for identification of systems having coloured noise*, Proceeding of Conference on Decision and Control, Hawaii, (1990).
- [40] R. SMITH AND M. D. (EDITORS), *The Modelling of Uncertainty in Control Systems*, Springer Verlag, 1994.

- [41] G. SZEGÖ, *Orthogonal Polynomials*, vol. 23, Colloquium publications - American Mathematical Society, New York, 1939.
- [42] S. TAKENAKA, *On the orthogonal functions and a new formula of interpolation*, Japanese Journal of Mathematics, II (1925), pp. 129–145.
- [43] T.SÖDERSTRÖM AND P.STOICA, *System Identification*, Prentice Hall, New York, 1989.
- [44] B. WAHLBERG, *System identification using Laguerre models*, IEEE Transactions on Automatic Control, AC-36 (1991), pp. 551–562.
- [45] B. WAHLBERG, *System identification using Kautz models*, IEEE Transactions on Automatic Control, AC-39 (1994), pp. 1276–1282.
- [46] B. WAHLBERG AND L.LJUNG, *Design variables for bias distribution in transfer function estimation*, IEEE Transactions on Automatic Control, AC-31 (1986), pp. 134–144.
- [47] H. WIDOM, *Studies in Real and Complex Analysis*, MAA Studies in Mathematics, Prentice Hall, Englewood Cliffs, NJ, 1965, ch. Toeplitz Matrices.
- [48] N. WIENER, *The Fourier Integral and Certain of its Applications*, Cambridge University Press, 1933.
- [49] Z. YUAN AND L. LJUNG, *Black box identification of multivariable transfer functions: Asymptotic properties and optimal input design*, International Journal of Control, 40 (1984), pp. 233–256.
- [50] Y. ZHU, *Black box identification of MIMO transfer functions: Asymptotic properties of prediction error models*, International Journal of Adaptive Control and Signal Processing, 3 (1989), pp. 357–373.
- [51] A. ZYGMUND, *Trigonometric Series*, Cambridge University Press, 1959.